

Autoregressive Processes

- The first-order autoregressive process, AR(1) is

$$y_t = \beta y_{t-1} + e_t$$

where e_t is $WN(0, \sigma^2)$

- Using the lag operator, we can write

$$(1 - \beta L)y_t = e_t$$

- If $\beta > 0$, y_{t-1} and y_t are positively correlated
- If $\beta < 0$, y_{t-1} and y_t are negatively correlated

Inversion

- By back-substitution

$$\begin{aligned}y_t &= \beta y_{t-1} + e_t \\ &= e_t + \beta(\beta y_{t-2} + e_{t-1}) \\ &= e_t + \beta e_{t-1} + \beta^2 e_{t-2} + \dots \\ &= \sum_{i=0}^{\infty} \beta^i e_{t-i}\end{aligned}$$

a general linear process with geometrically declining coefficients

- This inversion requires that $|\beta| < 1$
- $|\beta| < 1$ is required for stationarity

Importance of $|\beta| < 1$

- If $\beta=1$ then

$$y_t = e_t + e_{t-1} + e_{t-2} + \dots$$

does not converge, so the sum is not defined.

Mean and Variance

- By the formula for the unconditional mean and variance of a general linear process

$$E(y_t) = E\left(\sum_{i=0}^{\infty} \beta^i e_{t-i}\right) = 0$$

$$\text{var}(y_t) = \text{var}\left(\sum_{i=0}^{\infty} \beta^i e_{t-i}\right)$$

$$= \left(\sum_{i=0}^{\infty} \beta^{2i}\right) \sigma^2$$

$$= \frac{\sigma^2}{1 - \beta^2}$$

Another Variance Calculation

- Take variance of both sides of

$$y_t = \beta y_{t-1} + e_t$$

- Thus

$$\begin{aligned}\text{var}(y_t) &= \text{var}(\beta y_{t-1} + e_t) \\ &= \text{var}(\beta y_{t-1}) + \text{var}(e_t) \\ &= \beta^2 \text{var}(y_{t-1}) + \sigma^2\end{aligned}$$

- If y is variance stationary, we solve and find

$$\text{var}(y_t) = \text{var}(y_{t-1}) = \frac{\sigma^2}{1 - \beta^2}$$

$$|\beta| < 1$$

- If $|\beta|=1$ then

$$\text{var}(y_t) = \frac{\sigma^2}{1 - \beta^2}$$

is infinite

$$|\beta| = 1$$

- We calculated that

$$\text{var}(y_t) = \beta^2 \text{var}(y_{t-1}) + \sigma^2$$

- When $|\beta| = 1$, then

$$\text{var}(y_t) = \text{var}(y_{t-1}) + \sigma^2 > \text{var}(y_{t-1})$$

so the variance is increasing with t

- $|\beta| = 1$ is inconsistent with variance stationarity.
- $|\beta| < 1$ is necessary for stationarity.

Random Walk

- An AR(1) with $\beta=1$ is known as a random walk or unit root process

$$y_t = y_{t-1} + e_t$$

- By back-substitution

$$y_t = y_0 + \sum_{i=0}^t e_{t-i}$$

- The past never disappears. Shocks have permanent effects

Unit Root

- The random walk is called a **unit root** process because the lag operator $1-L$ has a “root” (intersection with the x-axis) at $L=1$
- It is called a **random walk** because it tends to wander without mean-reversion.
- If y_t is an AR(1) with a unit root ($\beta=1$) then its first difference $\Delta y_t = y_t - y_{t-1}$ is white noise

Conditional Mean and Variance of AR(1)

- Conditional mean:

$$E(y_t | \Omega_{t-1}) = E(\beta y_{t-1} + e_t | \Omega_{t-1}) = \beta y_{t-1}$$

- Conditional variance:

$$\begin{aligned} \text{var}(y_t | \Omega_{t-1}) &= \text{var}(y_t - E(y_t | \Omega_{t-1}) | \Omega_{t-1}) \\ &= \text{var}(e_t | \Omega_{t-1}) \\ &= \sigma^2 \end{aligned}$$

Autocovariance of AR(1)

- Take the equation

$$y_t = \beta y_{t-1} + e_t$$

- And then multiply both sides by y_{t-k}

$$y_{t-k} y_t = \beta y_{t-k} y_{t-1} + y_{t-k} e_t$$

- Then take expectations. Since e_t is white noise, it is uncorrelated with

$$E(y_{t-k} y_t) = \beta E(y_{t-k} y_{t-1}) + E(y_{t-k} e_t)$$

or

$$\gamma(k) = \beta \gamma(k-1)$$

Autocorrelation of AR(1)

- Dividing by the variance, this implies

$$\rho(k) = \beta\rho(k-1)$$

- We know

$$\rho(0) = 1$$

- Then

$$\rho(1) = \beta\rho(0) = \beta$$

$$\rho(2) = \beta\rho(1) = \beta^2$$

⋮

$$\rho(k) = \beta^k$$

Autocorrelation of AR(1)

- We have derived

$$\rho(k) = \beta^k$$

- The autocorrelation of the stationary AR(1) is a simple geometric decay ($|\beta| < 1$)
- If β is small, the autocorrelations decay rapidly to zero with k
- If β is large (close to 1) then the autocorrelations decay moderately
- The AR(1) parameter describes the persistence in the time series

One-Step-Ahead Forecast

- As we showed earlier

$$E(y_t | \Omega_{t-1}) = \beta y_{t-1}$$

- Thus

$$E(y_{T+1} | \Omega_T) = \beta y_T$$

- The optimal one-step-ahead forecast is a linear function of the final observed value

2-step-ahead forecast

- By back-substitution

$$\begin{aligned}y_t &= \beta y_{t-1} + e_t \\ &= e_t + \beta(\beta y_{t-2} + e_{t-1}) \\ &= \beta^2 y_{t-2} + e_t + \beta e_{t-1}\end{aligned}$$

- Thus

$$\begin{aligned}E(y_t | \Omega_{t-2}) &= E(\beta^2 y_{t-2} + e_t + \beta e_{t-1} | \Omega_{t-2}) \\ &= \beta^2 y_{t-2}\end{aligned}$$

- and

$$E(y_{T+2} | \Omega_T) = \beta^2 y_T$$

2-step-ahead forecast

- This shows that the optimal 2-step-ahead forecast is also a linear function of the final observed value, but with the coefficient β^2 .

$$E(y_{T+2} | \Omega_T) = \beta^2 y_T$$

h-step-ahead forecast

- Similarly

$$y_t = \beta^h y_{t-h} + e_t + \beta e_{t-1} + \cdots + \beta^{h-1} e_{t-h+1}$$

- So

$$\begin{aligned} E(y_t | \Omega_{t-h}) &= E(\beta^h y_{t-h} + e_t + \beta e_{t-1} + \cdots + \beta^{h-1} e_{t-h+1} | \Omega_{t-h}) \\ &= \beta^h y_{t-h} \end{aligned}$$

- Optimal forecast:

$$E(y_{T+h} | \Omega_T) = \beta^h y_T$$

Inversion of AR(1)

- By inverting the lag operator

$$(1 - \beta L)y_t = e_t$$

$$y_t = (1 - \beta L)^{-1} e_t$$

$$= \left(\sum_{i=0}^{\infty} \beta^i L^i \right) e_t$$

$$= \sum_{i=0}^{\infty} \beta^i e_{t-i}$$

- Which is the same as found by back substitution

Condition for Invertibility

- The operator $(1-\beta L)$ is invertible when $|\beta| < 1$
- This is the same as for the MA(1) model
- β is the inverse of the root of the polynomial $1-\beta L$
- The root of a function is the value where it crosses the x-axis
- The root of $1-\beta L$ is $1/\beta$, the inverse of the root is β
- Invertibility requires that the inverse of the root be less than one

AR(1) with Intercept

- An AR(1) with intercept is

$$y_t = \alpha + \beta y_{t-1} + e_t$$

Taking expectations

$$E(y_t) = \alpha + \beta E(y_{t-1}) + E(e_t)$$

- Thus

$$\mu = \alpha + \beta\mu$$

- and

$$\mu = \frac{\alpha}{1 - \beta}$$

Best Linear Predictor

- A linear predictor of y_t given y_{t-1} is

$$\alpha + \beta y_{t-1}$$

- The forecast error is

$$e_t = y_t - \alpha - \beta y_{t-1}$$

- The linear predictor which minimizes the expected squared forecast error solves

$$\min_{\alpha, \beta} E(y_t - \alpha - \beta y_{t-1})^2$$

Least-Squares

- The estimate of the expected squared linear forecast error is the sum of squared errors
- The least squares estimate

$$y_t = \hat{\alpha} + \hat{\beta}y_{t-1} + \hat{e}_t$$

minimizes the sum of squared errors, so is the estimate of the best linear predictor

- This is a linear regression, treating y_{t-1} as a regressor.

Unemployment Rate

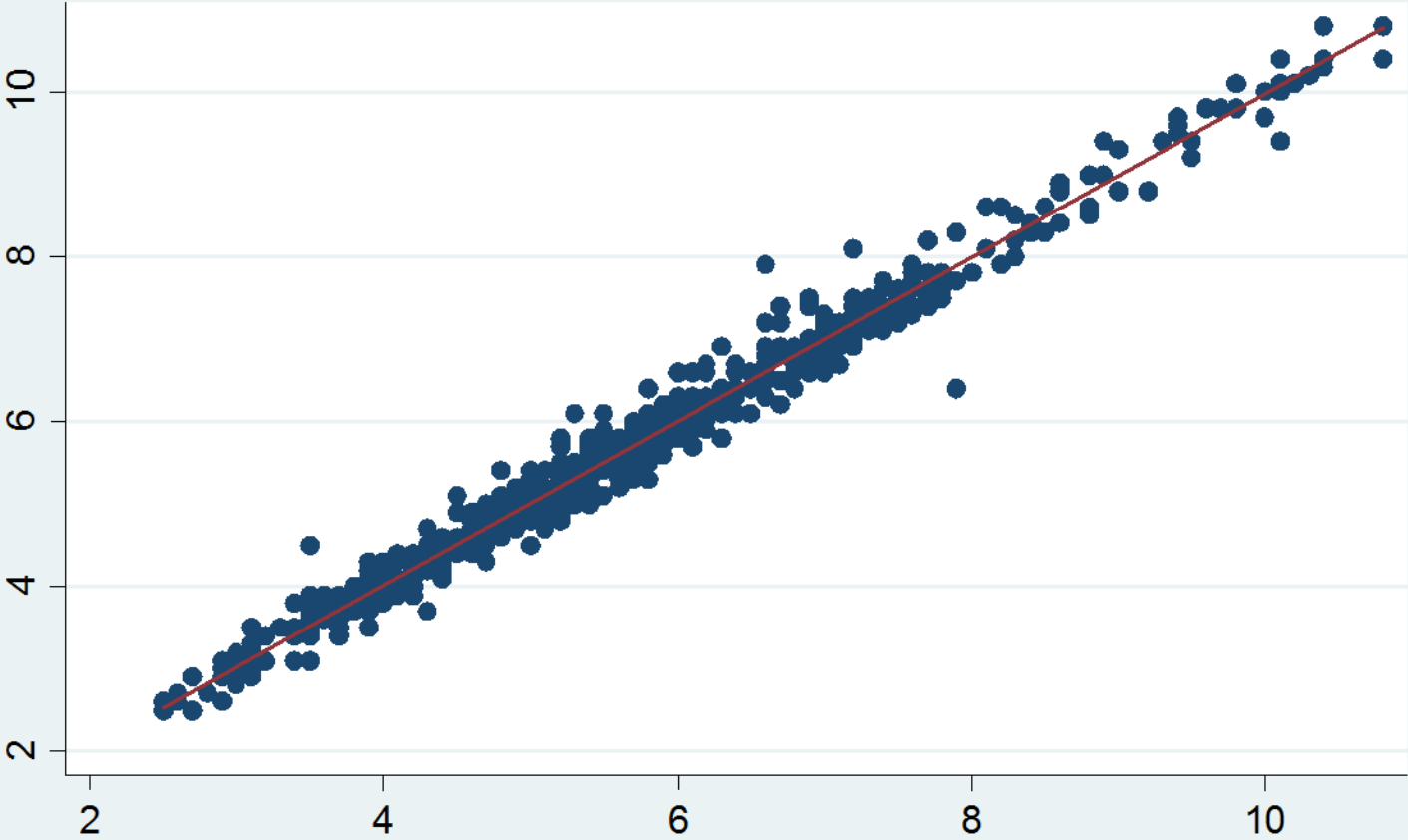
```
. regress ur L.ur
```

Source	SS	df	MS
Model	1775.33245	1	1775.33245
Residual	34.7193756	742	.046791611
Total	1810.05182	743	2.43613974

```
Number of obs =      744
F( 1, 742) =37941.25
Prob > F      = 0.0000
R-squared     = 0.9808
Adj R-squared = 0.9808
Root MSE     = .21631
```

ur	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ur L1.	.9934454	.0051002	194.79	0.000	.9834329	1.003458
_cons	.045538	.0299153	1.52	0.128	-.0131907	.1042667

Unemployment Rate



● ur — Fitted values

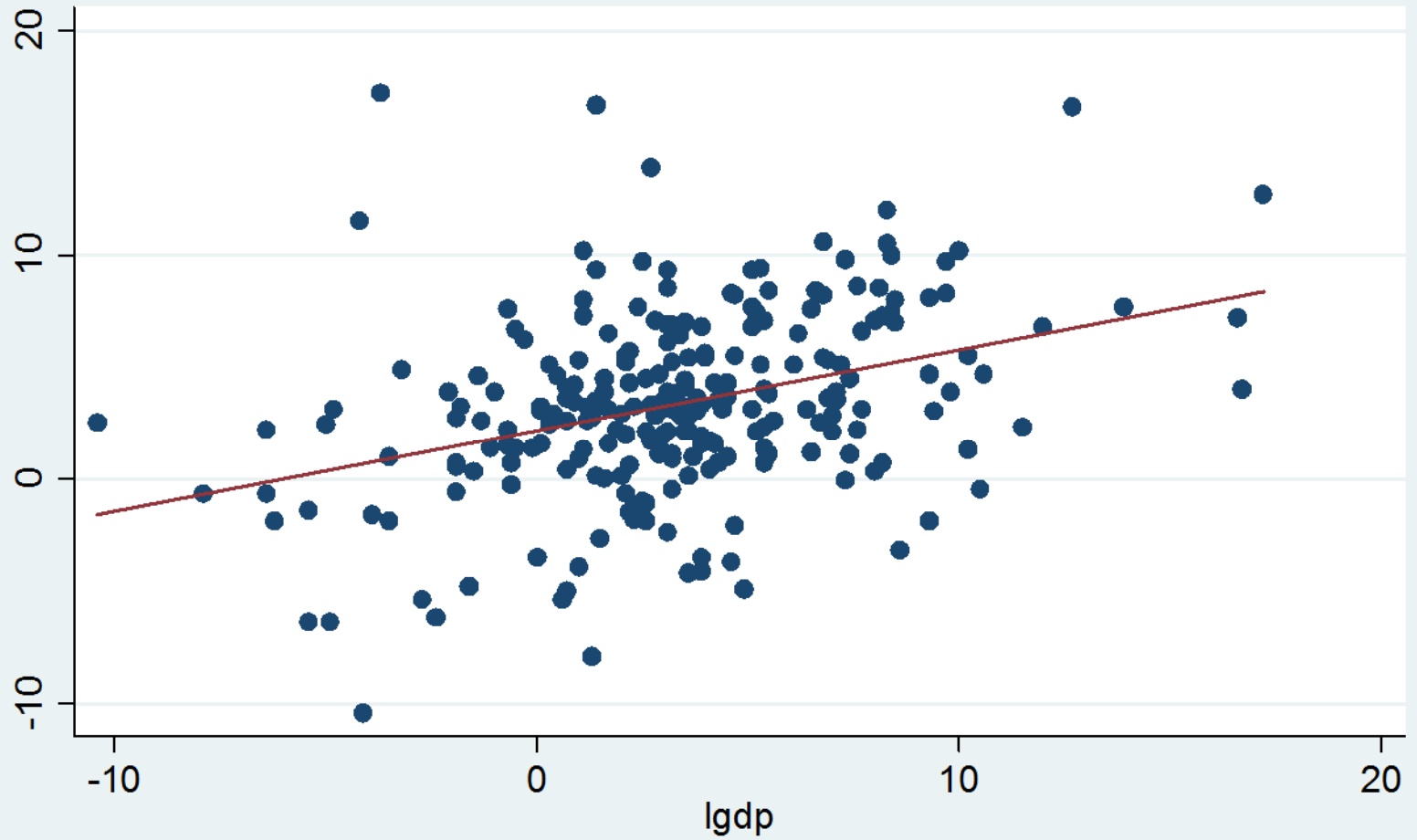
GDP Growth Rates

```
. reg gdp L.gdp
```

Source	SS	df	MS		
Model	548.684959	1	548.684959	Number of obs =	250
Residual	3662.76711	248	14.7692222	F(1, 248) =	37.15
Total	4211.45207	249	16.9134621	Prob > F =	0.0000
				R-squared =	0.1303
				Adj R-squared =	0.1268
				Root MSE =	3.8431

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gdp L1.	.3605283	.0591503	6.10	0.000	.2440274	.4770292
_cons	2.146711	.3123873	6.87	0.000	1.531441	2.761982

GDP Growth Rates



● GDP growth — Fitted values

One-Step-Ahead Forecast

- The optimal forecast for $T+1$ given T is

$$\hat{y}_{T+1|T} = \alpha + \beta y_T$$

- The forecast using the estimates is

$$\hat{y}_{T+1|T} = \hat{\alpha} + \hat{\beta} y_T$$

Example – Unemployment Rate

- The estimates were

ur						
L1.		.9934454	.0051002	194.79	0.000	.9834329
_cons		.045538	.0299153	1.52	0.128	-.0131907

$$y_t = 0.0455 + 0.993y_{t-1} + \hat{e}_t$$

- The value for Jan 2010 is 9.7%, so

$$\hat{y}_{2010:2} = 0.0455 + 0.993 \times 9.7 = 9.68$$

	y_t	y_{t-1}	fitted value
Jan-09	7.7	7.4	7.40
Feb-09	8.2	7.7	7.70
Mar-09	8.6	8.2	8.19
Apr-09	8.9	8.6	8.59
May-09	9.4	8.9	8.89
Jun-09	9.5	9.4	9.38
Jul-09	9.4	9.5	9.48
Aug-09	9.7	9.4	9.38
Sep-09	9.8	9.7	9.68
Oct-09	10.1	9.8	9.78
Nov-09	10	10.1	10.08
Dec-09	10	10	9.98
Jan-10	9.7	10	9.98
Feb-10	?	9.7	9.68

Example – GDP Growth

- The estimates were

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
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$$y_t = 2.14 + 0.361y_{t-1} + \hat{e}_t$$

- The value for 4th quarter 2009 is 5.7%, so

$$\hat{y}_{2010:1} = 2.14 + 0.361 \times 5.7 = 4.2\%$$

GDP Growth

	y_t	y_{t-1}	fitted
2008q4	-5.4	-2.7	1.2
2009q1	-6.4	-5.4	0.2
2009q2	-0.7	-6.4	-0.2
2009q3	2.2	-0.7	1.9
2009q4	5.7	2.2	2.9
2010q1	?	5.7	4.2

One-Step-Ahead Forecast Error

- The forecast error is

$$\begin{aligned}y_{T+1} - \hat{y}_{T+1|T} &= \alpha + \beta y_T + e_{T+1} - (\alpha + \beta y_T) \\ &= e_{T+1}\end{aligned}$$

- The forecast variance is

$$\text{var}(y_{T+1} - \hat{y}_{T+1|T}) = \text{var}(e_{T+1}) = \sigma^2$$

Forecast variance estimation

- Average of squared residuals

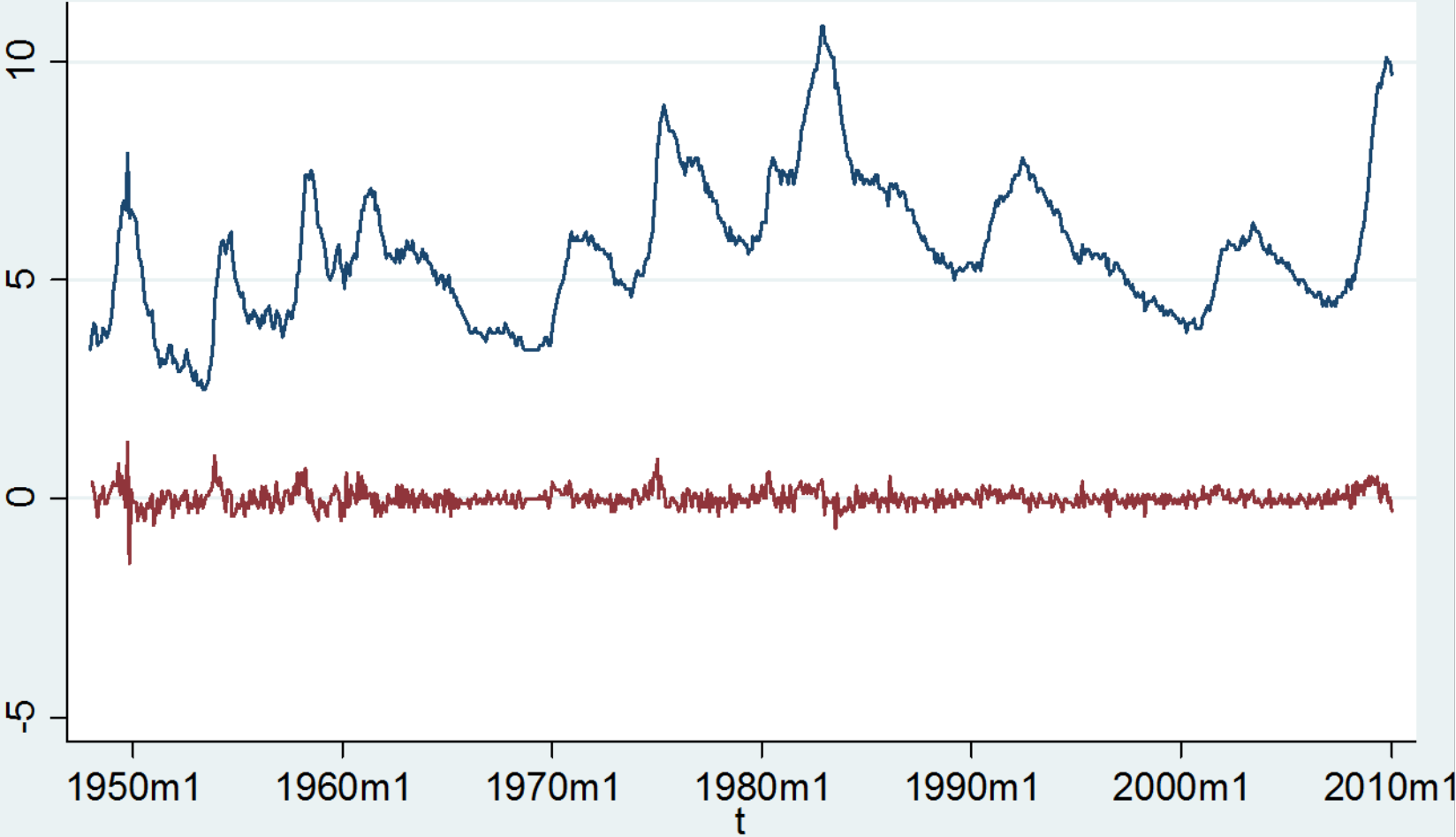
$$\hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T \hat{e}_t^2$$

$$\hat{\sigma} = \sqrt{\hat{\sigma}^2}$$

where the least-squares residuals are

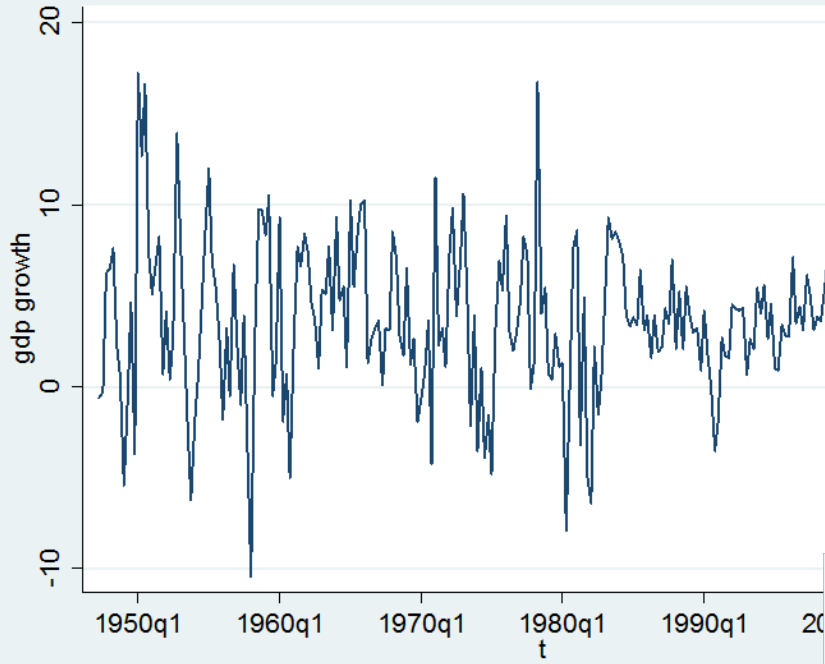
$$\hat{e}_t = y_t - \hat{\alpha} - \hat{\beta}y_{t-1}$$

Unemployment Rate

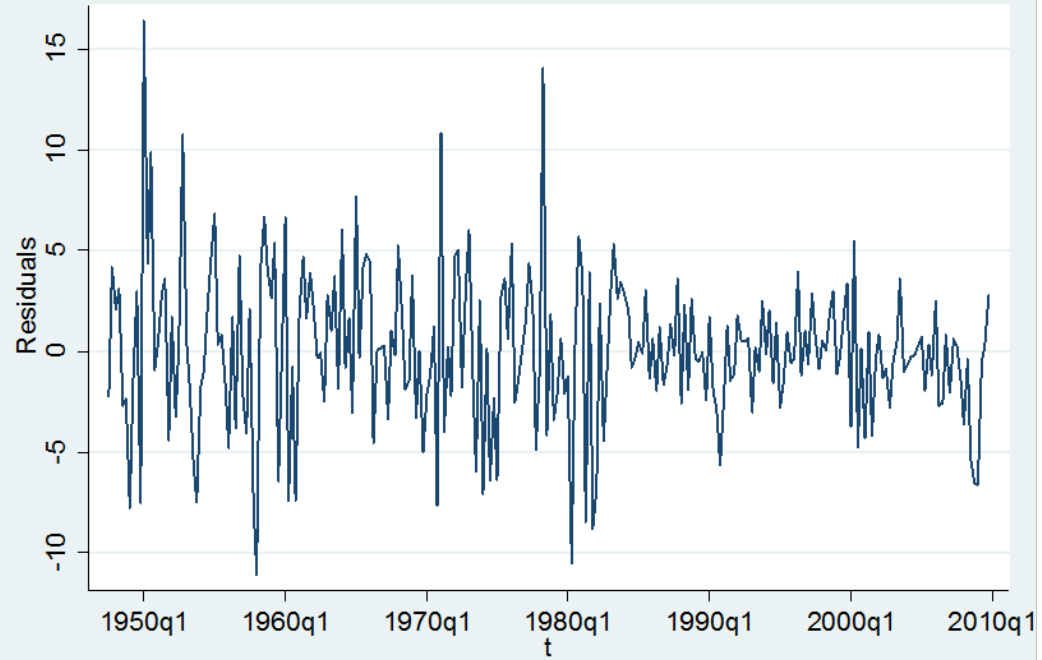


— ur — Residuals

GDP Growth



Residuals



One-Step-Ahead Intervals

- Normal Method

- Assume forecast error is normally distributed
- Forecast interval is point estimate, plus and minus the estimated standard deviation multiplied by a normal quantile

- For a 95% interval:

$$\hat{y}_{T+1|T} \pm \hat{\sigma} \cdot z_{.025} = \hat{y}_{T+1|T} \pm \hat{\sigma} \cdot 1.96$$

- For a 90% interval

$$\hat{y}_{T+1|T} \pm \hat{\sigma} \cdot z_{.05} = \hat{y}_{T+1|T} \pm \hat{\sigma} \cdot 1.645$$

Estimating Forecast Variance

- The estimated variance is 16.9
- The estimated st. dev. is 3.84

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Standard Dev Estimation

- In STATA, the estimate of σ or “root mean squared error” is saved after you estimate a regression in “e(rmse)”
- Better, the forecast standard deviation is stdf, through the command

predict s, stdf

Forecast Interval Construction

```
. tsappend, add(1)

. predict p if t>tq(2009q4)
(option xb assumed; fitted values)
(251 missing values generated)

. gen p1=p-1.645*e(rmse)
(251 missing values generated)

. gen p2=p+1.645*e(rmse)
(251 missing values generated)
```

- Point estimate = 4.2%
- 90% Interval = [-2.1%, 10.5%]