

Distribution of Least-Squares

- In classic regression, if the errors are iid normal, and independent of the regressors, then the least-squares estimates have an **exact** normal distribution, not just asymptotic
- This is not true in most time-series regressions.

Non-Classical Distributions

- Estimates in autoregressive models
 - Biased downwards
 - Skewed
 - Thick tails
- Especially
 - When autoregressive coefficients are large
 - Sample sizes are small
- These issues diminish in large samples

Example

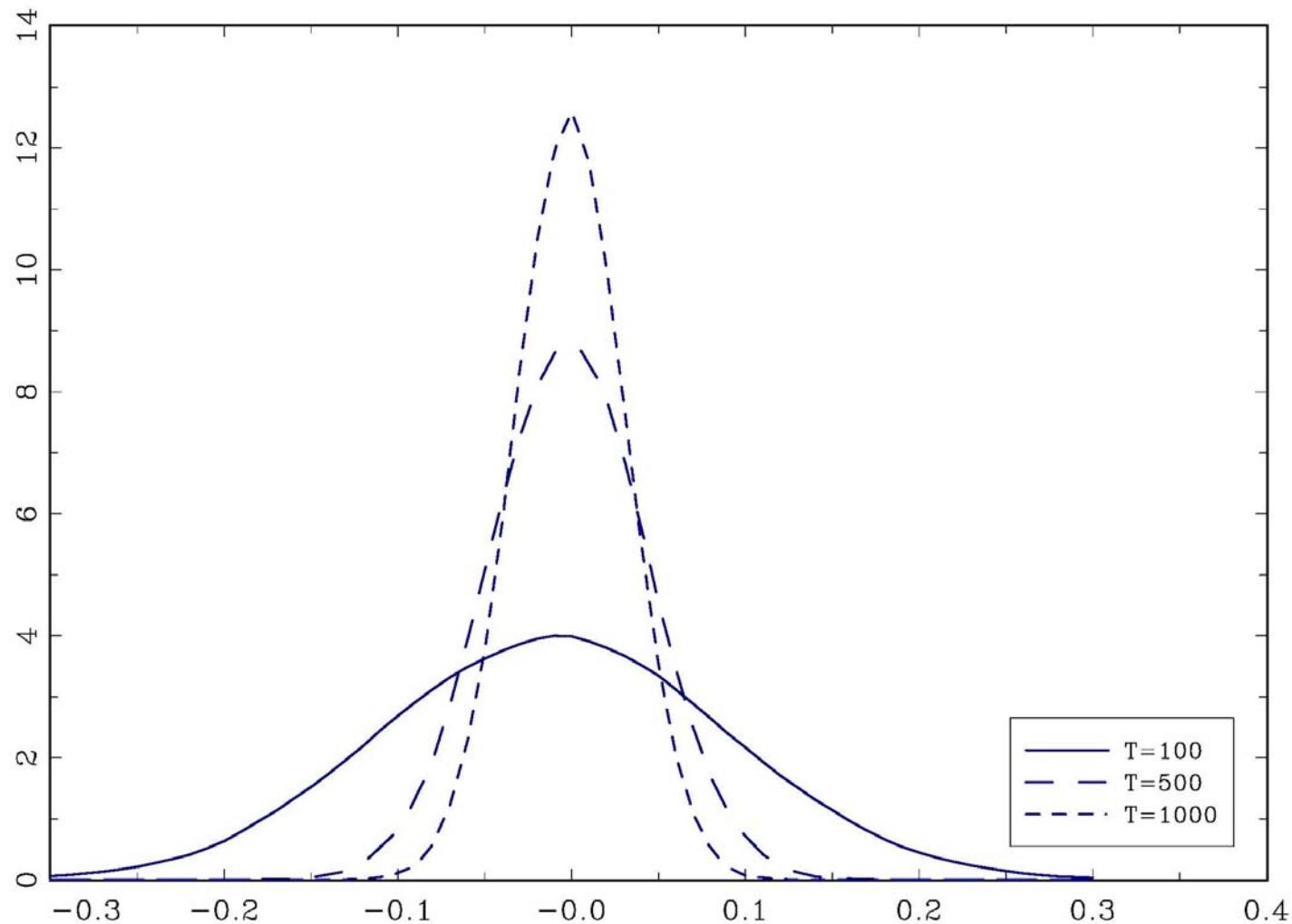
- Take the AR(1) model with intercept

$$y_t = \alpha + \beta y_{t-1} + e_t$$

- $e_t \sim N(0,1)$
- T=100, 500, 1000
- $\beta=0.0, \beta=0.5, \beta=0.9,$
- Numerically calculate distribution of least-squares estimate of β

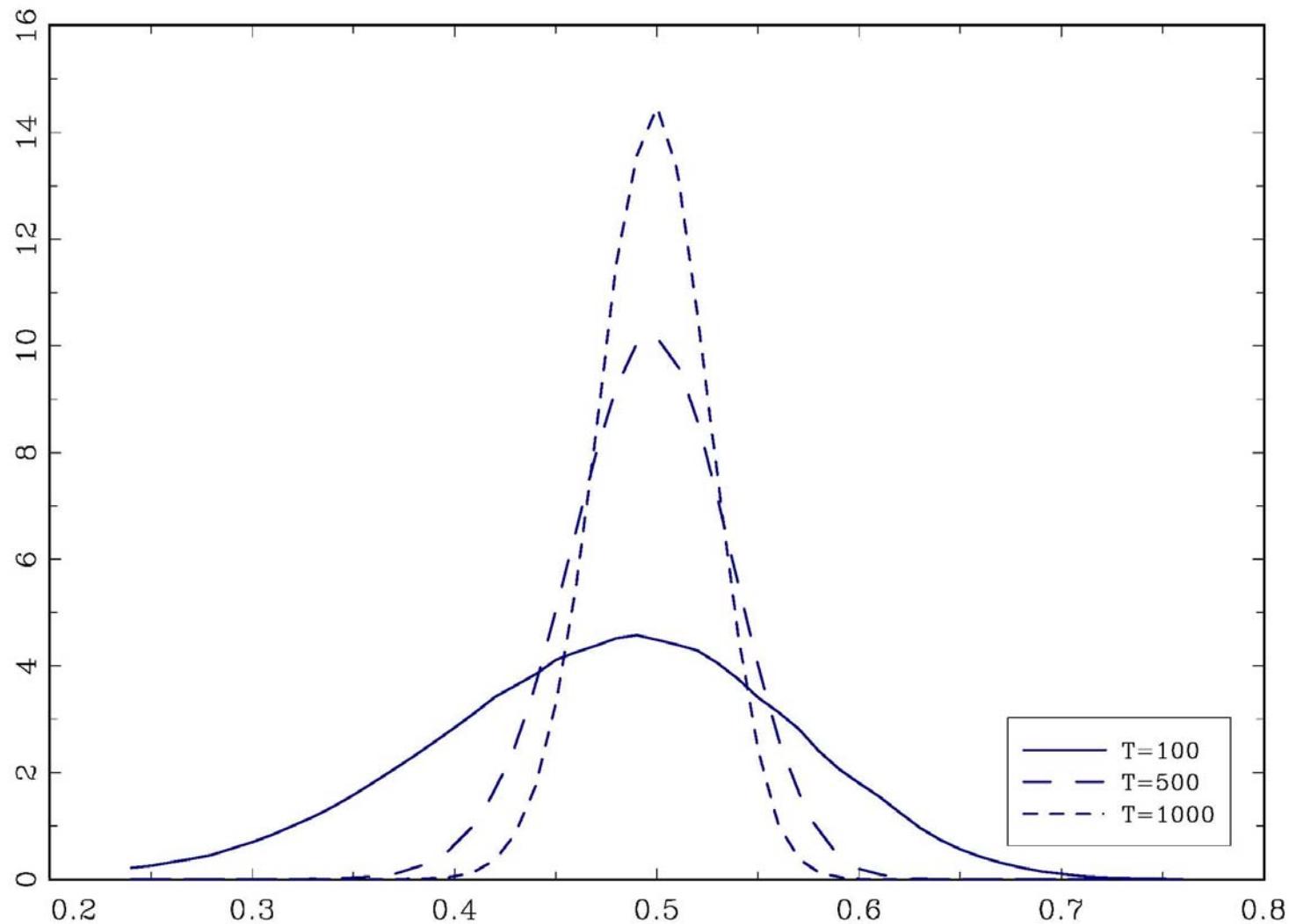
Distribution, $\beta=0.0$

$$\beta = 0$$



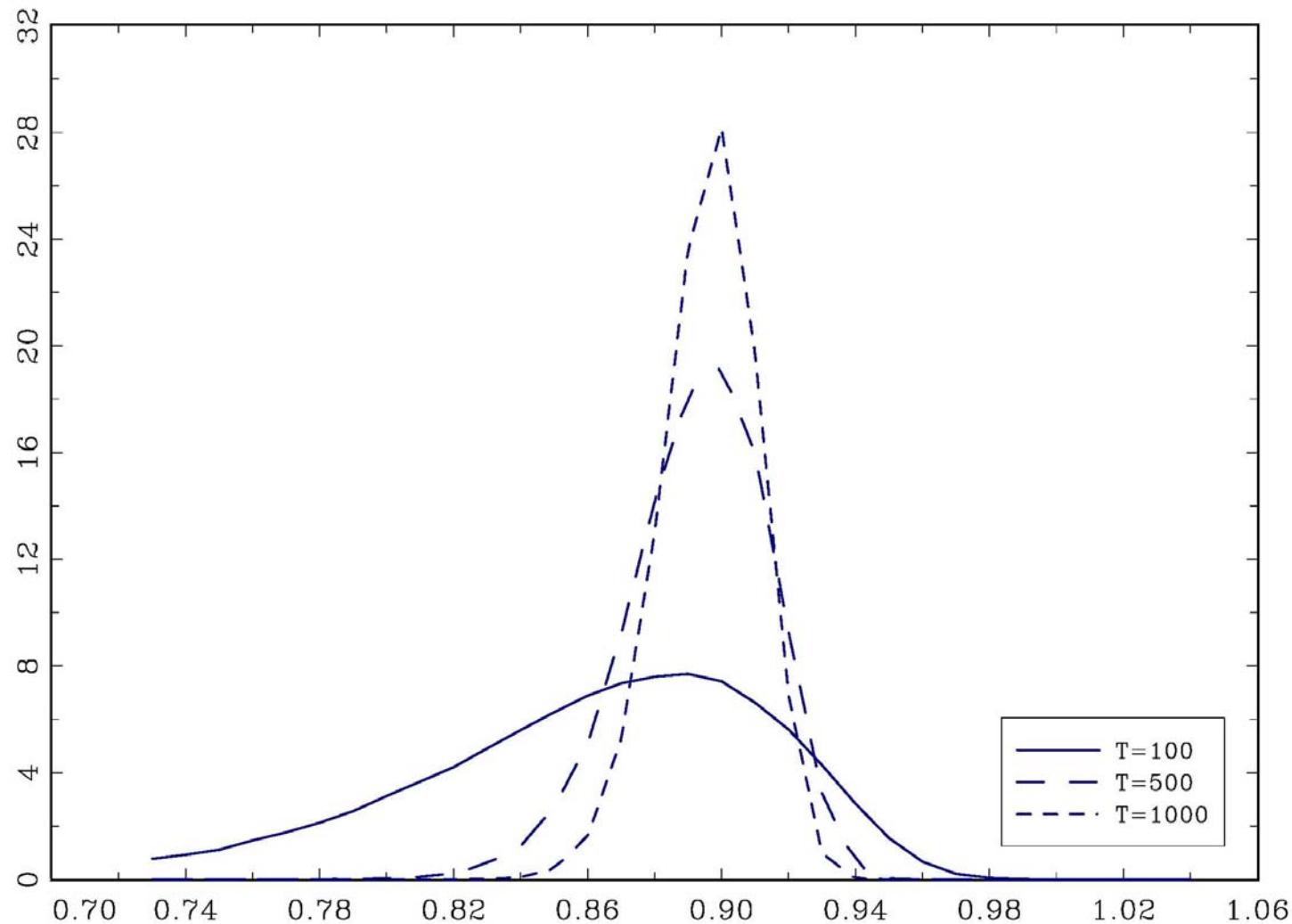
Distribution, $\beta=0.5$

$$\beta = 0.5$$



Distribution, $\beta=0.9$

$$\beta = 0.9$$



Interpretation

- Estimates of autoregressive parameters are random
- Even if regression error is normal, the parameter estimates are not normally distributed
- Distributions are less normal when AR coefficient is large
- Distributions are more concentrated and normal when sample size is large

Asymptotic Standard Deviation

- The least-squares estimate is asymptotically (approximately) normally distributed
- In the simple model $y_t = \beta x_t + e_t$

then

$$\hat{\beta} \stackrel{a}{\sim} N(\beta, \sigma_{\hat{\beta}}^2)$$

$$\sigma_{\hat{\beta}}^2 = \frac{1}{T} \frac{\text{var}(x_t e_t)}{\text{var}(x_t)^2}$$

- The standard deviation measures the precision of the estimate, but it is unknown.

Standard Errors

- Estimates of the standard deviations are called **standard errors**, and are reported in regression output
- They are used to measure estimation precision.

Classical standard errors

A **classic standard error** is an estimate of the standard deviation from the formula

$$\hat{\sigma}_{\hat{\beta}}^2 = \frac{1}{n} \frac{\text{var}(e_t)}{\text{var}(x_t)}$$

This formula is valid under conditional homoskedasticity

$$E(e_t | \Omega_{t-1}) = 0$$

$$E(e_t^2 | \Omega_{t-1}) = \sigma^2$$

Robust standard errors

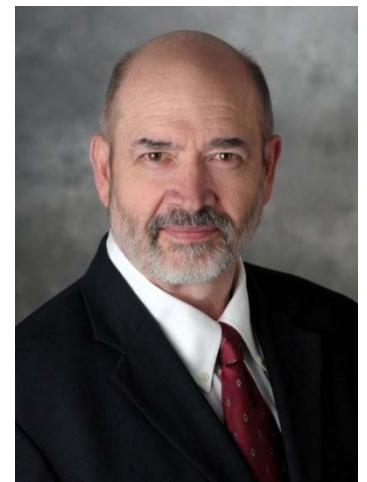
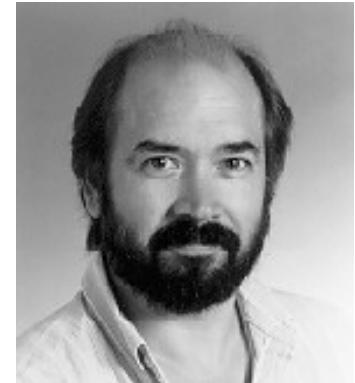
- “Robust” standard errors are estimates of

$$\sigma_{\beta} = \sqrt{\frac{1}{n} \frac{\text{var}(x_t e_t)}{[\text{var}(x_t)]^2}}$$

- These are the conventional standard errors for regression analysis
- Also known as “White” standard errors

Halbert White

- Professor Hal White, UCSD
- Leading contributor to econometric methods, especially time series analysis
- Introduced robust standard errors into econometrics (1980)
 - Most referenced paper in economics
- Founded Bates-White consulting firm, a leader in economic policy analysis



Have you seen robust standard errors?

- If you took an econometrics course other than 410, you may not be familiar with robust standard errors
- If you are currently taking 410, you won't cover robust standard errors until later in the course
 - Wooldridge uses the homoskedasticity assumption in the early part of his text
- Stock-Watson use robust standard errors throughout

Does the Choice Matter?

- Classic standard errors are for the assumption of conditional homoskedasticity

$$E(e_t^2 | \Omega_{t-1}) = \sigma^2$$

- This is **unforecastability in the variance**
 - This is not implied by conventional unforecastability
 - It may be a convenient approximation for macro data
 - It is a bad assumption (quite false) in financial data

Example: Stock Returns, AR(1)

r	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
L1. ^r	-.0163895	.0179105	-0.92	0.360	-.051507 .0187279
_cons	.0013626	.0003728	3.66	0.000	.0006317 .0020935

r	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
L1. ^r	-.0163895	.032377	-0.51	0.613	-.0798718 .0470927
_cons	.0013626	.0003878	3.51	0.000	.0006022 .002123

- The robust standard error on the AR(1) coefficient is almost twice as large as the conventional standard error

Computation

- In STATA, the default is conventional standard errors.
- They are automatically reported with the regress (reg) command
- For robust standard errors, use the “r” option

.reg y x, r

Example: Real GDP Growth

. reg gdp L(1/4).gdp

Source	SS	df	MS	Number of obs	=	247
Model	662.232234	4	165.558059	F(4, 242)	=	11.39
Residual	3518.78213	242	14.540422	Prob > F	=	0.0000
Total	4181.01437	246	16.9959934	R-squared	=	0.1584

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
gdp					
L1.	.327656	.0640344	5.12	0.000	.2015202 .4537919
L2.	-.1466135	.0670302	2.19	0.030	-.0145764 .2786506
L3.	-.0980287	.066934	-1.46	0.144	-.2298764 .0338189
L4.	-.0889209	.0644466	-1.38	0.169	-.2158689 .038027
_cons	2.378427	.389677	6.10	0.000	1.610836 3.146019

With Robust st. errors

```
. reg gdp L(1/4).gdp, r
```

Linear regression

Number of obs = 247
F(4, 242) = 8.85
Prob > F = 0.0000
R-squared = 0.1584
Root MSE = 3.8132

gdp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
L1.	.327656	.076895	4.26	0.000	.1761871 .479125
L2.	.1466135	.0858808	1.71	0.089	-.0225558 .3157828
L3.	-.0980287	.0728951	-1.34	0.180	-.2416186 .0455611
L4.	-.0889209	.0790354	-1.13	0.262	-.244606 .0667641
_cons	2.378427	.4731312	5.03	0.000	1.446447 3.310408

Robust st. errors

- With the “r” option

.reg y x, r

- You get robust
 - Standard errors
 - t statistics and p-values
 - test statistics

Annoyance

- In STATA, with the “r” option, STATA omits sum of squared error table
 - Yet this can be useful
- So both commands may be useful

.reg y x

.reg y x, r

Interpretation of standard errors

- The standard errors measure precision of the estimate
 - Forecasts use *estimated* coefficients.
- Small standard errors mean the estimate is precise
 - Good for forecasting
- Large standard errors mean the estimate is not precise
 - Bad for forecasting
 - Inaccurate estimates leads to inaccurate forecasts

		Robust Std. Err.	t	P> t	[95% Conf. Interval]	
gdp	Coef.					
gdp						
L1.	.327656	.076895	4.26	0.000	.1761871	.479125
L2.	.1466135	.0858808	1.71	0.089	-.0225558	.3157828
L3.	-.0980287	.0728951	-1.34	0.180	-.2416186	.0455611
L4.	-.0889209	.0790354	-1.13	0.262	-.244606	.0667641
_cons	2.378427	.4731312	5.03	0.000	1.446447	3.310408

Interpretation of t-statistics

- “t” is the coefficient estimate divided by the standard error.
- It is used to *test* if the coefficient is zero
 - “P>|t|” is the p-value of the t-statistic
 - If p<.05 you “reject” the hypothesis of a zero coefficient
- Hypothesis tests are useful for assessing economic theories
 - But are less useful for picking good forecasting models

gdp	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
gdp					
L1.	.327656	.076895	4.26	0.000	.1761871 .479125
L2.	.1466135	.0858808	1.71	0.089	-.0225558 .3157828
L3.	-.0980287	.0728951	-1.34	0.180	-.2416186 .0455611
L4.	-.0889209	.0790354	-1.13	0.262	-.244606 .0667641
_cons	2.378427	.4731312	5.03	0.000	1.446447 3.310408

Interpretation of Confidence Interval

- The 95% interval is the coefficient estimate plus and minus 1.96 times the standard error
- Helps gauge possible values for the true coefficient
- Useful tool

gdp	Coef.	Robust Std. Err.	t	P> t 	[95% Conf. Interval]	
gdp						
L1.	.327656	.076895	4.26	0.000	.1761871	.479125
L2.	.1466135	.0858808	1.71	0.089	-.0225558	.3157828
L3.	-.0980287	.0728951	-1.34	0.180	-.2416186	.0455611
L4.	-.0889209	.0790354	-1.13	0.262	-.244606	.0667641
_cons	2.378427	.4731312	5.03	0.000	1.446447	3.310408

Summary

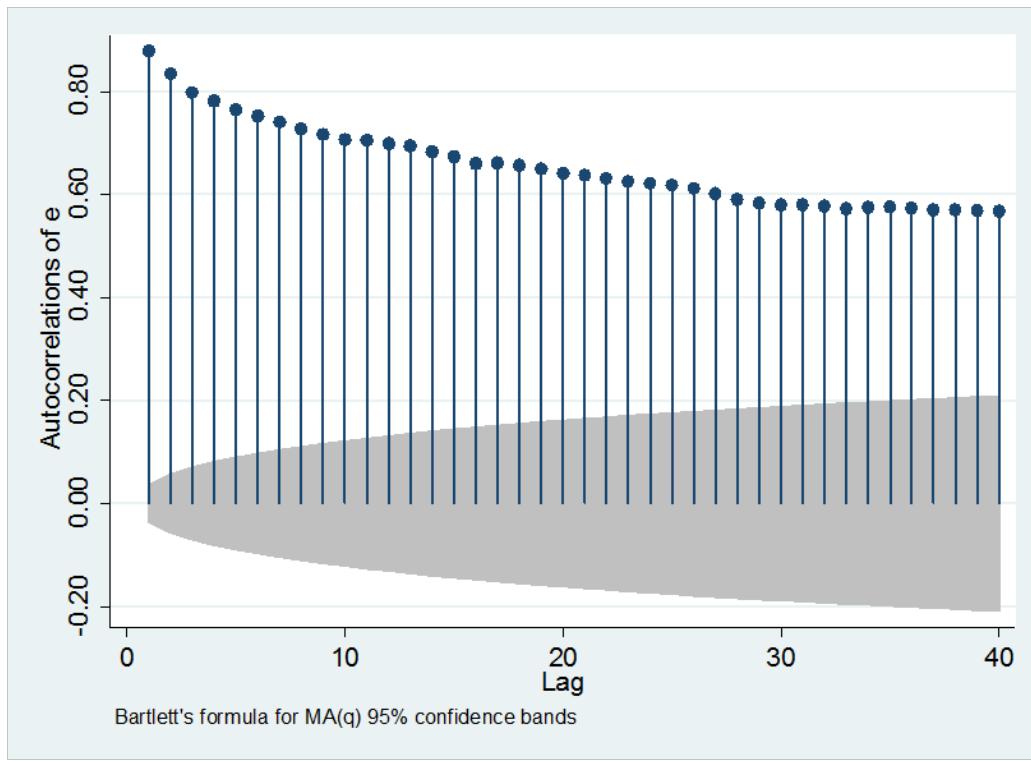
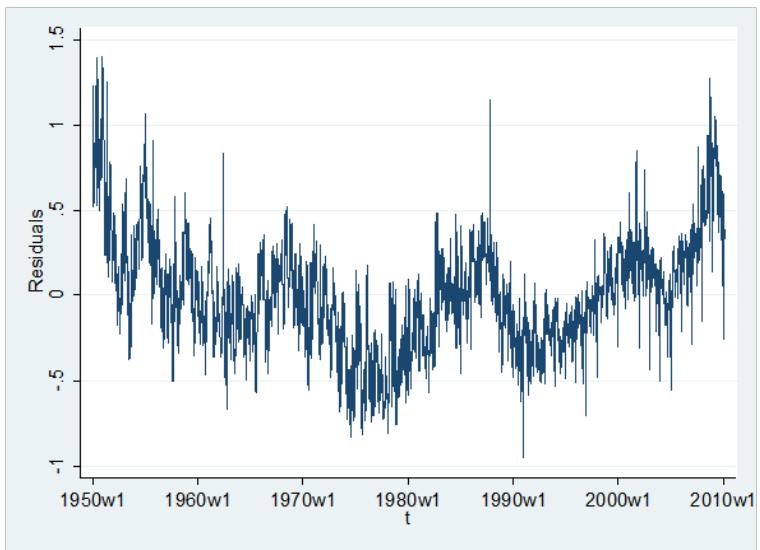
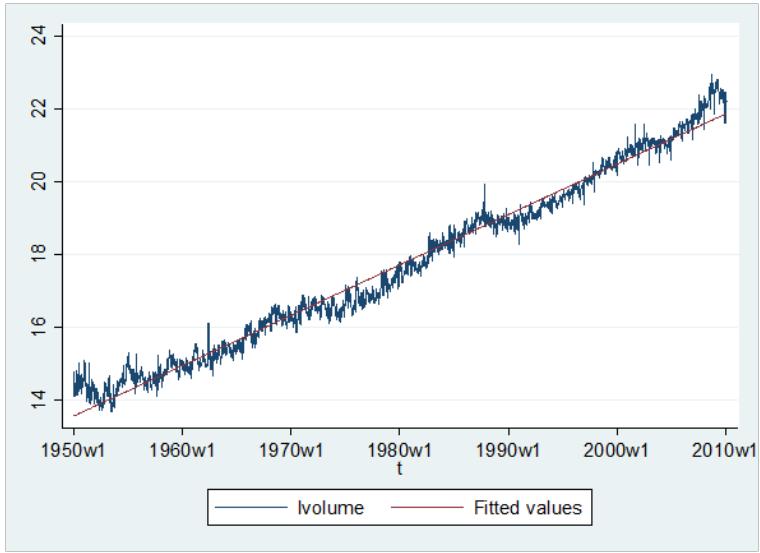
- In one-step-ahead forecast regressions with unforecastable errors
 - Robust standard errors generally appropriate
 - Classical standard errors appropriate under conditional homoskedasticity

Regression with Correlated Errors

$$y_t = \alpha + \beta x_t + e_t$$

- In some regression models, the errors are correlated
 - Pure Trend Models
 - Pure Seasonality Models
- In these models the errors can be correlated
- Classical and robust standard errors are not appropriate

Example: Stock Volume



Least-Squares Variance Formula

Recall for $v_t = x_t e_t$

$$\text{var}(\hat{\beta}) \sim \frac{\text{var}\left(\sum_{t=1}^T v_t\right)}{[T \text{var}(x_t)]^2}$$

When the v are uncorrelated

$$\text{var}\left(\sum_{t=1}^T v_t\right) = \sum_{t=1}^T \text{var}(v_t) = T \text{var}(v_t)$$

$$\text{var}(\hat{\beta}) \sim \frac{\text{var}(v_t)}{T[\text{var}(x_t)]^2}$$

General Formula

Define

$$f_T = \frac{\text{var}\left(\sum_{t=1}^T v_t\right)}{T \text{ var}(v_t)}$$

When the v are uncorrelated $f_T=1$, otherwise not.

Then

$$\text{var}(\hat{\beta}) \sim \frac{\text{var}(x_t e_t)}{T[\text{var}(x_t)]^2} f_T$$

Adjustment Factor

- The asymptotic variance of least-squares is the conventional, multiplied by an adjustment factor for the serial correlation

$$\text{var}(\hat{\beta}) \sim \frac{\text{var}(x_t e_t)}{T[\text{var}(x_t)]^2} f_T$$

Autocovariance of v

- We want a useful formula for

$$f_T = \frac{\text{var}\left(\sum_{t=1}^T v_t\right)}{T \text{ var}(v_t)}$$

- Since $E(v_t) = 0$, then

$$E(v_t^2) = \text{var}(v_t)$$

$$E(v_t v_j) = \text{cov}(v_t v_j) = \gamma(t - j)$$

the autocovariance of v_t

Variance of sum of correlated v

$$\begin{aligned}\text{var}\left(\sum_{t=1}^T v_t\right) &= E\left(\sum_{t=1}^T v_t\right)^2 \\ &= E\left(\sum_{t=1}^T v_t \sum_{j=1}^T v_j\right) \\ &= \sum_{t=1}^T \sum_{j=1}^T E(v_t v_j) \\ &= \sum_{t=1}^T \sum_{j=1}^T \gamma(t-j)\end{aligned}$$

Adjustment Factor

$$f_T = \frac{\text{var}\left(\sum_{t=1}^T v_t\right)}{T \text{ var}(v_t)} = \frac{1}{T} \sum_{t=1}^T \sum_{j=1}^T \rho(t-j)$$

- Where the $\rho(t-j)$ are the autocorrelations of v_t

- This double sum is the sum of all the elements in the matrix

$$\begin{bmatrix} \rho(0) & \rho(1) & \rho(2) & \cdots & \rho(T-1) \\ \rho(1) & \rho(0) & \rho(1) & \cdots & \rho(T-2) \\ \rho(2) & \rho(1) & \rho(0) & \cdots & \rho(T-3) \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \rho(T-1) & \rho(T-2) & \rho(T-3) & \cdots & \rho(0) \end{bmatrix}$$

- There are
 - T of the $\rho(0)$
 - $2(T-1)$ of the $\rho(1)$
 - $2(T-2)$ of the $\rho(2)$
 - ...

$$T + \sum_{j=1}^{T-1} 2(T-j)\rho(j)$$

Adjustment Factor

- Dividing by T

$$\begin{aligned}f_T &= \frac{1}{T} \sum_{t=1}^T \sum_{j=1}^T \rho(t-j) \\&= 1 + \sum_{j=1}^{T-1} 2\left(\frac{T-j}{T}\right) \rho(j)\end{aligned}$$

- If T is large

$$f_T \rightarrow 1 + 2 \sum_{j=1}^{\infty} \rho(j) = f$$

Summary: Least-Squares Variance

- When the errors are correlated

$$\text{var}(\hat{\beta}) \sim \frac{\text{var}(x_t e_t)}{T[\text{var}(x_t)]^2} f$$

$$f = 1 + 2 \sum_{j=1}^{\infty} \rho(j)$$

- The conventional formula is multiplied by an adjustment for autocorrelation

HAC Estimation

- Estimation of f
 - For variances and standard errors under autocorrelation
- Called heteroskedasticity and autocorrelation consistent (HAC) variance estimation
- Multiply conventional variance estimates by estimates of f

HAC Estimation

- The adjustment is

$$f = 1 + 2 \sum_{j=1}^{\infty} \rho(j)$$

where $\rho(j)$ are the autocorrelations of $v_t = x_t e_t$

- Estimate $\rho(j)$ by sample autocorrelations using least-squares residuals
- But in a sample of length T we cannot estimate all autocorrelations well

Unweighted HAC Estimator

- For some **truncation parameter** m ,

$$\hat{f} = 1 + 2 \sum_{j=1}^m \hat{\rho}(j)$$

- Original proposal
 - L. Hansen, Hodrick (1978)
 - Hal White (1982)
- Difficiencies
 - This estimator is not smooth in the truncation parameter
 - The sample estimate can be negative

Lars Hansen

- Professor Lars Hansen, U Chicago
- Invented Generalized Method of Moments, the leading estimation method for applied econometrics
- Introduced unweighted HAC estimator for multi-step regression models
- On short list for Nobel in economics

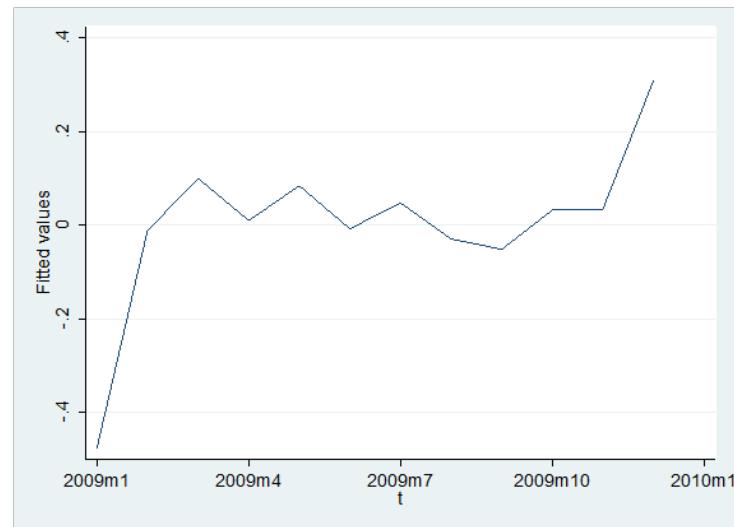
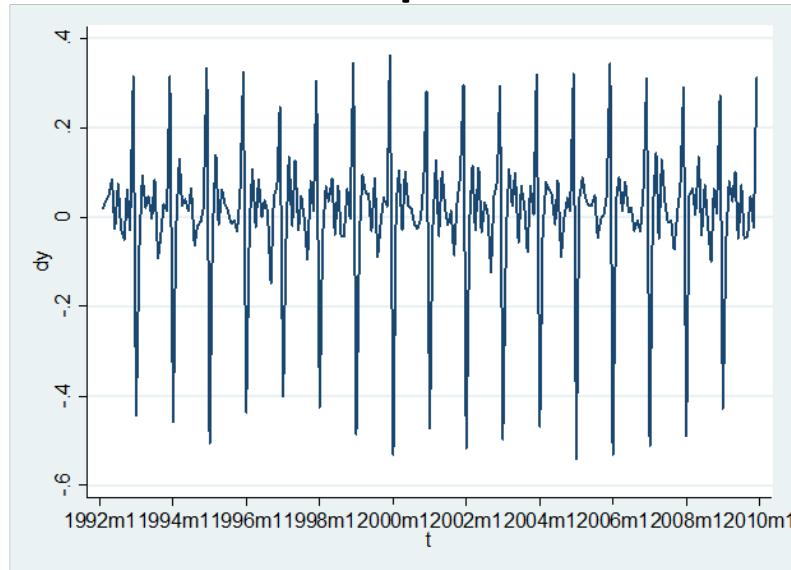


Example of Negative Estimate

- Take $m=1$
- Then $\hat{f} = 1 + 2\hat{\rho}(1) < 0$
if estimated $\rho(1) < -1/2$

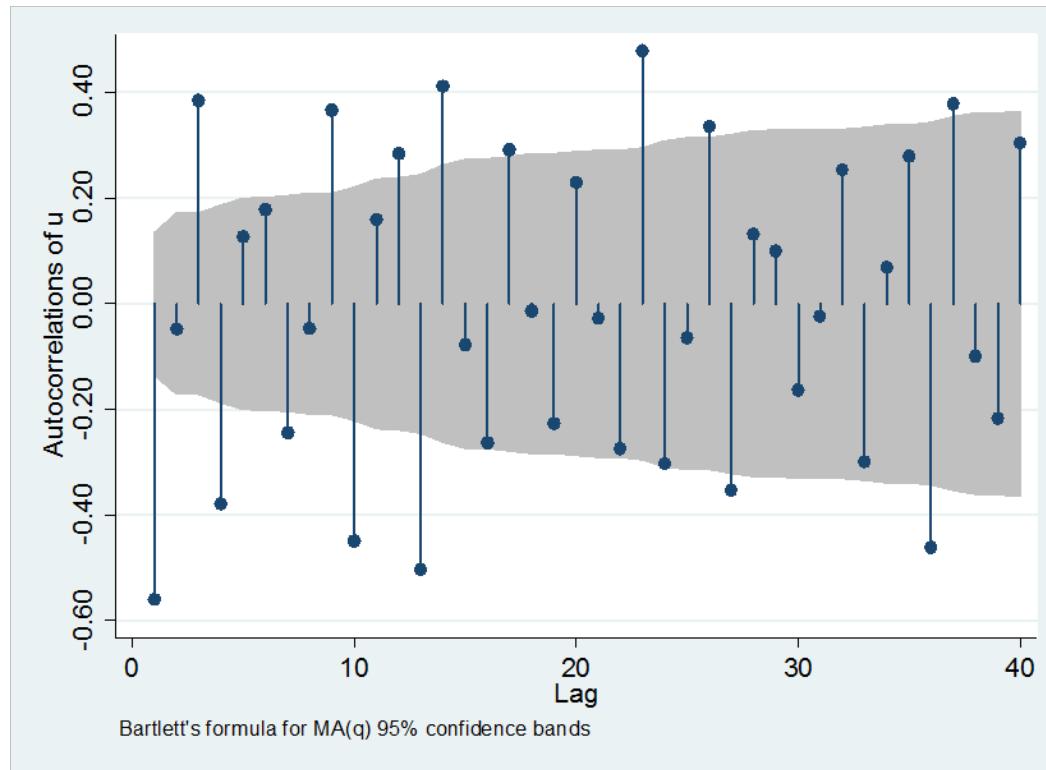
Example: Liquor Sales

- Transform to growth rates
- Monthly change in log liquor sales
- Regress on Seasonal Dummies only to obtain seasonal pattern



Autocorrelation of Residual

- The first autocorrelation is less than $-1/2$



Weighted HAC Estimator

$$\hat{f} = 1 + 2 \sum_{j=1}^m \left(\frac{m-j}{m} \right) \hat{\rho}(j)$$

- Called Newey-West variance estimator
 - Whitney Newey, Ken West (1987)
- This weighted estimator is always positive
- Smoothly changes in truncation parameter m

Whitney Newey and Ken West

- Professor Whitney Newey, MIT
 - Leading econometric theorist
- Professor Ken West, Wisconsin
 - Macroeconomist, econometrician
 - Forecast evaluation and comparison
- Joint paper in 1987
 - Weighted HAC estimator
 - One of the most referenced papers in econometrics



Computation

- In STATA, replace **regress** command with **newey** command

.newey y x, lag(m)

- You supply the truncation parameter “m”
- Similar to regression with robust standard errors
- These are identical

.newey y x, lag(0)

.reg y x, r

Example: Liquor Sales

. reg dy b12.m,r

Linear regression

Number of obs = 215
F(11, 203) = 423.80
Prob > F = 0.0000
R-squared = 0.9613
Root MSE = .0347

dy	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]
m					
1	-.788371	.0120765	-65.28	0.000	-.8121825 -.7645595
2	-.3218705	.0105478	-30.52	0.000	-.3426677 -.3010733
3	-.2103181	.0094619	-22.23	0.000	-.2289744 -.1916619
4	-.3002915	.010514	-28.56	0.000	-.3210222 -.2795607
5	-.2258118	.0100036	-22.57	0.000	-.245536 -.2060876
6	-.3185358	.0096047	-33.16	0.000	-.3374735 -.2995981
7	-.2618824	.0100737	-26.00	0.000	-.2817449 -.2420198
8	-.3392591	.0107775	-31.48	0.000	-.3605093 -.3180088
9	-.3624475	.0123023	-29.46	0.000	-.3867042 -.3381907
10	-.2782956	.010299	-27.02	0.000	-.2986023 -.257989
11	-.2761872	.0108553	-25.44	0.000	-.2975908 -.2547835
_cons	.3099733	.0065735	47.16	0.000	.2970122 .3229343

With Newey-West standard errors

. newey dy b12.m, lag(12)

Regression with Newey-West standard errors
maximum lag: 12

Number of obs = 215
F(11, 203) = 908.34
Prob > F = 0.0000

dy	Coef.	Newey-West Std. Err.	t	P> t	[95% Conf. Interval]
m					
1	-.788371	.0149943	-52.58	0.000	-.8179356 -.7588064
2	-.3218705	.0093479	-34.43	0.000	-.3403018 -.3034391
3	-.2103181	.0100234	-20.98	0.000	-.2300816 -.1905547
4	-.3002915	.0087418	-34.35	0.000	-.3175278 -.2830551
5	-.2258118	.0128307	-17.60	0.000	-.2511104 -.2005132
6	-.3185358	.0087245	-36.51	0.000	-.335738 -.3013336
7	-.2618824	.0090442	-28.96	0.000	-.279715 -.2440498
8	-.3392591	.0134996	-25.13	0.000	-.3658765 -.3126416
9	-.3624475	.0075171	-48.22	0.000	-.377269 -.3476259
10	-.2782956	.0116472	-23.89	0.000	-.3012606 -.2553307
11	-.2761872	.0126533	-21.83	0.000	-.3011359 -.2512384
_cons	.3099733	.0066381	46.70	0.000	.2968848 .3230618

Truncation Parameter

- m should be large when autocorrelation is large
- Sophistical data-dependent methods to pick m have been developed, but are not in STATA
- Stock-Watson default (explanatory x's)

$$m = 0.75T^{1/3}$$

- Trend/Seasonal default

$$m = 1.4T^{1/3}$$

Derivation of Defaults

- Due to Andrews (1991)
- The optimal m minimizes the mean-squared error of the estimate of f
- When v_t is an AR(1) with coefficient ρ ,
Andrews found the optimal m is

$$m = CT^{1/3}$$

$$C = \left(\frac{6\rho^2}{(1-\rho^2)^2} \right)^{1/3}$$

Donald Andrews

- Professor Donald Andrews, Yale
- Leading econometric theorist
- Contributions to time-series
 - Optimal selection of truncation parameter
 - Tests for structural change



Default Values

$$m = CT^{1/3}$$

$$C = \left(\frac{6\rho^2}{(1-\rho^2)^2} \right)^{1/3}$$

- Stock-Watson
 - If both x_t and e_t are AR(1) with coef $\frac{1}{2}$, then $v_t=x_te_t$ has AR(1) coefficient $\rho=.25$. Plug this in, and $C=.75$
- Trend-Seasonal
 - If x_t is trend and/or seasonal and e_t are AR(1) with coef $\frac{1}{2}$, then $v_t=x_te_t$ has AR(1) coefficient $\rho=.5$. Plug this in, and $C=1.4$

Liquor Sales again

. dis 1.4*e(N)^(1/3)
8.387017

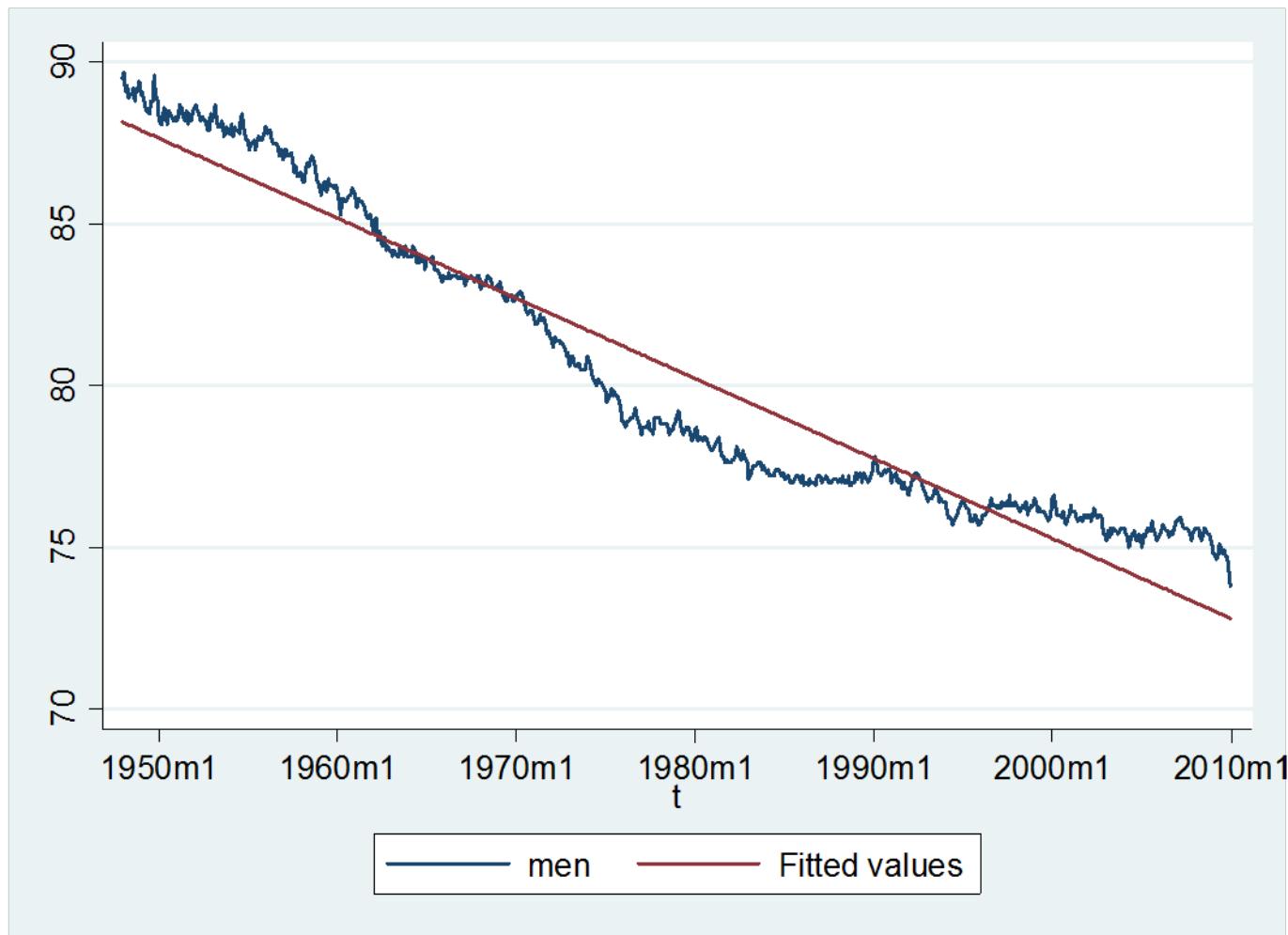
. newey dy b12.m, lag(8)

Regression with Newey-West standard errors
maximum lag: 8

Number of obs = 215
F(11, 203) = 736.19
Prob > F = 0.0000

dy	Coef.	Newey-West Std. Err.	t	P> t	[95% Conf. Interval]
m					
1	-.788371	.0146673	-53.75	0.000	-.8172907 -.7594513
2	-.3218705	.0089781	-35.85	0.000	-.3395727 -.3041682
3	-.2103181	.0097191	-21.64	0.000	-.2294815 -.1911548
4	-.3002915	.0097151	-30.91	0.000	-.319447 -.281136
5	-.2258118	.0116748	-19.34	0.000	-.2488312 -.2027924
6	-.3185358	.0089588	-35.56	0.000	-.3362001 -.3008715
7	-.2618824	.00916	-28.59	0.000	-.2799433 -.2438214
8	-.3392591	.0126319	-26.86	0.000	-.3641655 -.3143526
9	-.3624475	.0091312	-39.69	0.000	-.3804516 -.3444434
10	-.2782956	.0106888	-26.04	0.000	-.2993709 -.2572204
11	-.2761872	.0126343	-21.86	0.000	-.3010984 -.2512759
_cons	.3099733	.0065735	47.16	0.000	.2970122 .3229343

Example: Men's Labor Force Participation Rate, Trend Model



```
. reg m t
```

Source	SS	df	MS	Number of obs	=	744
Model	14659.2499	1	14659.2499	F(1, 742)	=	9554.38
Residual	1138.4477	742	1.53429609	Prob > F	=	0.0000
Total	15797.6976	743	21.2620426	R-squared	=	0.9279

men	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
t	-.0206675	.0002114	-97.75	0.000	-.0210826	-.0202524
_cons	85.18169	.0661519	1287.67	0.000	85.05182	85.31156

```
. dis 1.4*e(N)^(1/3)  
12.685834
```

```
. newey men t, lag(13)
```

Regression with Newey-West standard errors
maximum lag: 13

Number of obs = 744
F(1, 742) = 692.69
Prob > F = 0.0000

men	Coef.	Newey-West		t	P> t	[95% Conf. Interval]	
		Std. Err.					
t	-.0206675	.0007853	-26.32	0.000	-.0222091	-.0191259	
_cons	85.18169	.2168636	392.79	0.000	84.75595	85.60743	

Summary

- In one-step-ahead forecast regressions
- If the errors are serially uncorrelated
 - Use Robust standard errors
 - reg with r option
- If the errors are correlated
 - Use Newey-West standard errors
 - newey y x, lag(m)
 - In pure trend or seasonality models
 - Set $m=1.4T^{1/3}$
 - In dynamic regression
 - Set $m=.75T^{1/3}$

h-step-ahead forecasts

- In the AR(1) Model

$$y_t = \alpha + \beta y_{t-1} + e_t$$

- The optimal h-step forecasting regression takes the form

$$y_t = \alpha + \beta^h y_{t-h} + u_t$$

$$u_t = e_t + \beta e_{t-1} + \beta^2 e_{t-2} + \cdots + \beta^{h-1} e_{t-h+1}$$

- The error u_t is a correlated MA(h-1)
 - Unless $\beta=0$

h -step-ahead models

- In any h -step model

$$y_t = \alpha + \beta y_{t-h} + u_t$$

the variable $v_t = y_{t-h} e_t$ is generally serially correlated

- Generally MA($h-1$)
- Correct adjustment term

$$f = 1 + 2 \sum_{j=1}^{h-1} \rho(j)$$

Newey-West Standard Errors

- Standard errors can be estimated using the Newey-West method
- Truncation parameter set to forecast horizon
 - $m=h$

$$\hat{f} = 1 + 2 \sum_{j=1}^{h-1} \left(\frac{h-j}{h} \right) \hat{\rho}(j)$$

Example: Unemployment Rate

- 12-month-ahead forecast with 4 AR lags
 - Robust standard errors:

```
. reg ur L(12/15).ur ,r
```

Linear regression

Number of obs = 730
F(4, 725) = 139.36
Prob > F = 0.0000
R-squared = 0.4955
Root MSE = 1.1088

ur	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
L12.	1.686434	.2920485	5.77	0.000	1.113072	2.259795
L13.	-.0698989	.3908098	-0.18	0.858	-.837153	.6973552
L14.	-.5401552	.3461042	-1.56	0.119	-1.219641	.1393309
L15.	-.4100512	.2538791	-1.62	0.107	-.9084772	.0883747
_cons	1.94875	.1705347	11.43	0.000	1.613949	2.28355

Example: Unemployment Rate

- Newey-West standard errors:
 - Standard errors on lag 13 and 14 decrease by half
 - Standard error on constant more than doubles
- `newey ur L(12/15).ur, lag(12)`

Regression with Newey-West standard errors
maximum lag: 12

Number of obs = 730
F(4, 725) = 21.00
Prob > F = 0.0000

ur	Coef.	Newey-West		t	P> t	[95% Conf. Interval]	
		Std. Err.					
ur	1.686434	.273372	6.17	0.000	1.149738	2.223129	
L12.	-.0698989	.1564772	-0.45	0.655	-.3771014	.2373036	
L13.	-.5401552	.1378278	-3.92	0.000	-.8107445	-.2695658	
L14.	-.4100512	.246517	-1.66	0.097	-.8940236	.0739212	
_cons	1.94875	.4550687	4.28	0.000	1.05534	2.842159	

newey and forecasting

- **predict** works after **newey** command
- **e(rmse)** does not work, only after **regress** or **reg**
 - rmse not computed or reported
- Use **newey** to assess model and examine coefficients
- Use **reg** to compute out-of-sample forecast intervals

Summary

- In one-step-ahead forecast regressions
 - If the errors are serially uncorrelated, use `r` option
 - If the errors are correlated
 - Use **newey** for standard errors
 - In pure trend or seasonality models set $m=1.4T^{1/3}$
 - In dynamic regression set $m=.75T^{1/3}n$
 - Use **reg** and **e(rmse)** for forecast intervals
- In h-step-ahead forecast regressions
 - Use **newey** with $m=h$ for standard errors
 - Use **reg** and **e(rmse)** for forecast intervals