

# Autoregressive Processes

- The first-order autoregressive process, AR(1) is

$$y_t = \beta y_{t-1} + e_t$$

where  $e_t$  is WN(0,  $\sigma^2$ )

- Using the lag operator, we can write

$$(1 - \beta L)y_t = e_t$$

- If  $\beta > 0$ ,  $y_{t-1}$  and  $y_t$  are positively correlated
- If  $\beta < 0$ ,  $y_{t-1}$  and  $y_t$  are negatively correlated

# Inversion

- By back-substitution

$$\begin{aligned}y_t &= \beta y_{t-1} + e_t \\ &= e_t + \beta(\beta y_{t-2} + e_{t-1}) \\ &= e_t + \beta e_{t-1} + \beta^2 e_{t-2} + \dots \\ &= \sum_{i=0}^{\infty} \beta^i e_{t-i}\end{aligned}$$

a general linear process with geometrically declining coefficients

- This inversion requires that  $|\beta| < 1$
- $|\beta| < 1$  is required for stationarity

# Importance of $|\beta| < 1$

- If  $\beta=1$  then

$$y_t = e_t + e_{t-1} + e_{t-2} + \dots$$

does not converge, so the sum is not defined.

# Mean and Variance

- By the formula for the unconditional mean and variance of a general linear process

$$E(y_t) = E\left(\sum_{i=0}^{\infty} \beta^i e_{t-i}\right) = 0$$

$$\text{var}(y_t) = \text{var}\left(\sum_{i=0}^{\infty} \beta^i e_{t-i}\right)$$

$$= \left(\sum_{i=0}^{\infty} \beta^{2i}\right) \sigma^2$$

$$= \frac{\sigma^2}{1 - \beta^2}$$

# Another Variance Calculation

- Take variance of both sides of

$$y_t = \beta y_{t-1} + e_t$$

- Thus

$$\begin{aligned}\text{var}(y_t) &= \text{var}(\beta y_{t-1} + e_t) \\ &= \text{var}(\beta y_{t-1}) + \text{var}(e_t) \\ &= \beta^2 \text{var}(y_{t-1}) + \sigma^2\end{aligned}$$

- If  $y$  is variance stationary, we solve and find

$$\text{var}(y_t) = \text{var}(y_{t-1}) = \frac{\sigma^2}{1 - \beta^2}$$

$$|\beta| < 1$$

- If  $|\beta|=1$  then

$$\text{var}(y_t) = \frac{\sigma^2}{1 - \beta^2}$$

is infinite

$$|\beta| = 1$$

- We calculated that

$$\text{var}(y_t) = \beta^2 \text{var}(y_{t-1}) + \sigma^2$$

- When  $|\beta| = 1$ , then

$$\text{var}(y_t) = \text{var}(y_{t-1}) + \sigma^2 > \text{var}(y_{t-1})$$

so the variance is increasing with  $t$

- $|\beta| = 1$  is inconsistent with variance stationarity.
- $|\beta| < 1$  is necessary for stationarity.

# Random Walk

- An AR(1) with  $\beta=1$  is known as a random walk or unit root process

$$y_t = y_{t-1} + e_t$$

- By back-substitution

$$y_t = y_0 + \sum_{i=0}^t e_{t-i}$$

- The past never disappears. Shocks have permanent effects



# Unit Root

- The random walk is called a **unit root** process because the lag operator  $1-L$  has a “root” (intersection with the x-axis) at  $L=1$
- It is called a **random walk** because it tends to wander without mean-reversion.
- If  $y_t$  is an AR(1) with a unit root ( $\beta=1$ ) then its first difference  $\Delta y_t = y_t - y_{t-1}$  is white noise

# Conditional Mean and Variance of AR(1)

- Conditional mean:

$$E(y_t | \Omega_{t-1}) = E(\beta y_{t-1} + e_t | \Omega_{t-1}) = \beta y_{t-1}$$

- Conditional variance:

$$\begin{aligned} \text{var}(y_t | \Omega_{t-1}) &= \text{var}(y_t - E(y_t | \Omega_{t-1}) | \Omega_{t-1}) \\ &= \text{var}(e_t | \Omega_{t-1}) \\ &= \sigma^2 \end{aligned}$$

# Autocovariance of AR(1)

- Take the equation

$$y_t = \beta y_{t-1} + e_t$$

- And then multiply both sides by  $y_{t-k}$

$$y_{t-k} y_t = \beta y_{t-k} y_{t-1} + y_{t-k} e_t$$

- Then take expectations. Since  $e_t$  is white noise, it is uncorrelated with

$$E(y_{t-k} y_t) = \beta E(y_{t-k} y_{t-1}) + E(y_{t-k} e_t)$$

or

$$\gamma(k) = \beta \gamma(k-1)$$

# Autocorrelation of AR(1)

- Dividing by the variance, this implies

$$\rho(k) = \beta\rho(k-1)$$

- We know

$$\rho(0) = 1$$

- Then

$$\rho(1) = \beta\rho(0) = \beta$$

$$\rho(2) = \beta\rho(1) = \beta^2$$

⋮

$$\rho(k) = \beta^k$$

# Autocorrelation of AR(1)

- We have derived

$$\rho(k) = \beta^k$$

- The autocorrelation of the stationary AR(1) is a simple geometric decay ( $|\beta| < 1$ )
- If  $\beta$  is small, the autocorrelations decay rapidly to zero with  $k$
- If  $\beta$  is large (close to 1) then the autocorrelations decay moderately
- The AR(1) parameter describes the persistence in the time series

# One-Step-Ahead Forecast

- As we showed earlier

$$E(y_t | \Omega_{t-1}) = \beta y_{t-1}$$

- Thus

$$E(y_{T+1} | \Omega_T) = \beta y_T$$

- The optimal one-step-ahead forecast is a linear function of the final observed value

# 2-step-ahead forecast

- By back-substitution

$$\begin{aligned}y_t &= \beta y_{t-1} + e_t \\ &= e_t + \beta(\beta y_{t-2} + e_{t-1}) \\ &= \beta^2 y_{t-2} + e_t + \beta e_{t-1}\end{aligned}$$

- Thus

$$\begin{aligned}E(y_t | \Omega_{t-2}) &= E(\beta^2 y_{t-2} + e_t + \beta e_{t-1} | \Omega_{t-2}) \\ &= \beta^2 y_{t-2}\end{aligned}$$

- and

$$E(y_{T+2} | \Omega_T) = \beta^2 y_T$$

# 2-step-ahead forecast

- This shows that the optimal 2-step-ahead forecast is also a linear function of the final observed value, but with the coefficient  $\beta^2$ .

$$E(y_{T+2} | \Omega_T) = \beta^2 y_T$$



# h-step-ahead forecast

- Similarly

$$y_t = \beta^h y_{t-h} + e_t + \beta e_{t-1} + \cdots + \beta^{h-1} e_{t-h+1}$$

- So

$$\begin{aligned} E(y_t | \Omega_{t-h}) &= E(\beta^h y_{t-h} + e_t + \beta e_{t-1} + \cdots + \beta^{h-1} e_{t-h+1} | \Omega_{t-h}) \\ &= \beta^h y_{t-h} \end{aligned}$$

- Optimal forecast:

$$E(y_{T+h} | \Omega_T) = \beta^h y_T$$

# Inversion of AR(1)

- By inverting the lag operator

$$(1 - \beta L)y_t = e_t$$

$$y_t = (1 - \beta L)^{-1} e_t$$

$$= \left( \sum_{i=0}^{\infty} \beta^i L^i \right) e_t$$

$$= \sum_{i=0}^{\infty} \beta^i e_{t-i}$$

- Which is the same as found by back substitution

# Condition for Invertibility

- The operator  $(1-\beta L)$  is invertible when  $|\beta| < 1$
- This is the same as for the MA(1) model
- $\beta$  is the inverse of the root of the polynomial  $1-\beta L$
- The root of a function is the value where it crosses the x-axis
- The root of  $1-\beta L$  is  $1/\beta$ , the inverse of the root is  $\beta$
- Invertibility requires that the inverse of the root be less than one

# AR(1) with Intercept

- An AR(1) with intercept is

$$y_t = \alpha + \beta y_{t-1} + e_t$$

Taking expectations

$$E(y_t) = \alpha + \beta E(y_{t-1}) + E(e_t)$$

- Thus

$$\mu = \alpha + \beta \mu$$

- and

$$\mu = \frac{\alpha}{1 - \beta}$$

# Best Linear Predictor

- A linear predictor of  $y_t$  given  $y_{t-1}$  is

$$\alpha + \beta y_{t-1}$$

- The forecast error is

$$e_t = y_t - \alpha - \beta y_{t-1}$$

- The linear predictor which minimizes the expected squared forecast error solves

$$\min_{\alpha, \beta} E(y_t - \alpha - \beta y_{t-1})^2$$

# Least-Squares

- The estimate of the expected squared linear forecast error is the sum of squared errors
- The least squares estimate

$$y_t = \hat{\alpha} + \hat{\beta}y_{t-1} + \hat{e}_t$$

minimizes the sum of squared errors, so is the estimate of the best linear predictor

- This is a linear regression, treating  $y_{t-1}$  as a regressor.

# Unemployment Rate

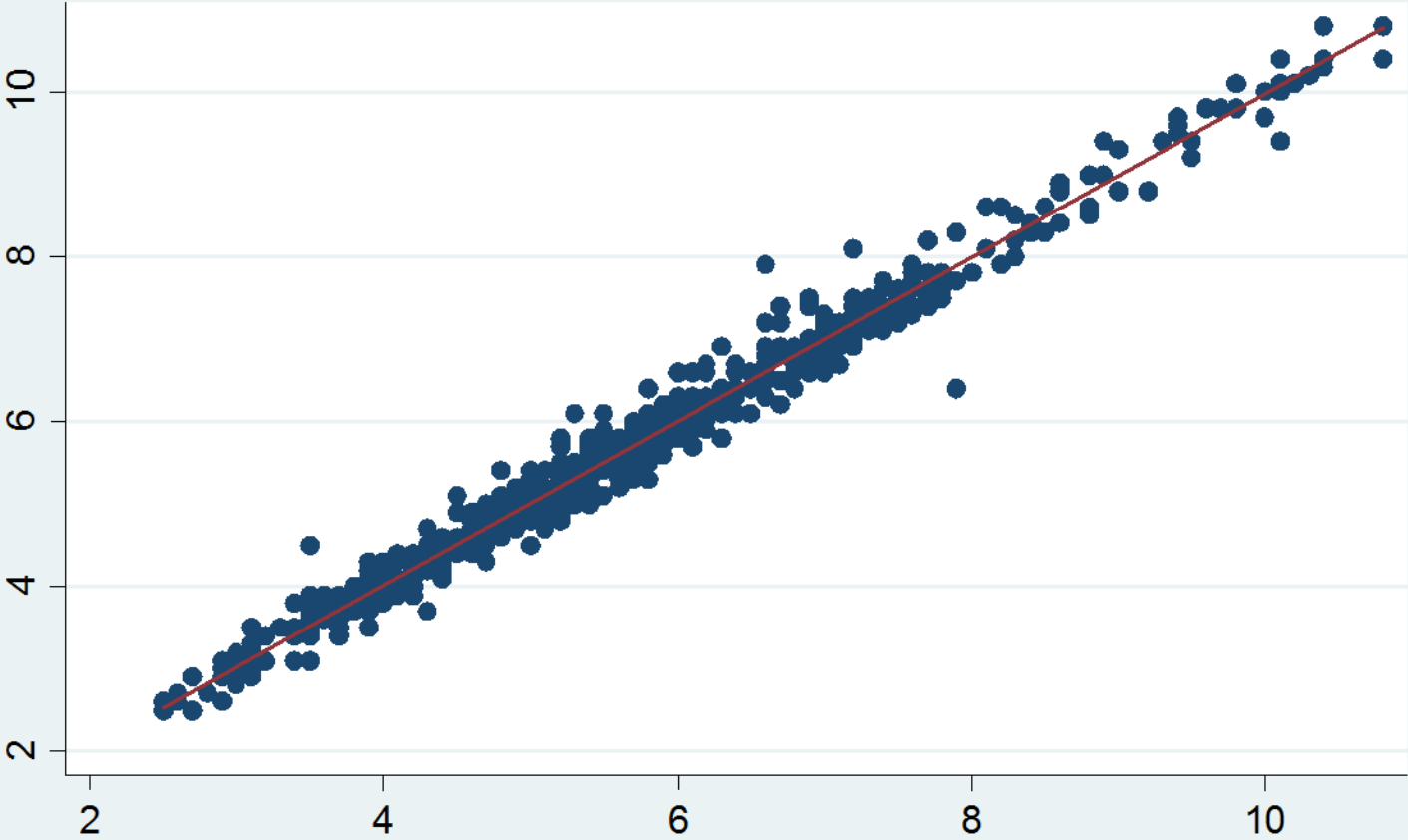
. regress ur L.ur

Source	SS	df	MS
Model	1775.33245	1	1775.33245
Residual	34.7193756	742	.046791611
Total	1810.05182	743	2.43613974

Number of obs = 744  
 F( 1, 742) = 37941.25  
 Prob > F = 0.0000  
 R-squared = 0.9808  
 Adj R-squared = 0.9808  
 Root MSE = .21631

ur	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
ur L1.	.9934454	.0051002	194.79	0.000	.9834329	1.003458
_cons	.045538	.0299153	1.52	0.128	-.0131907	.1042667

# Unemployment Rate



● ur — Fitted values



# GDP Growth Rates

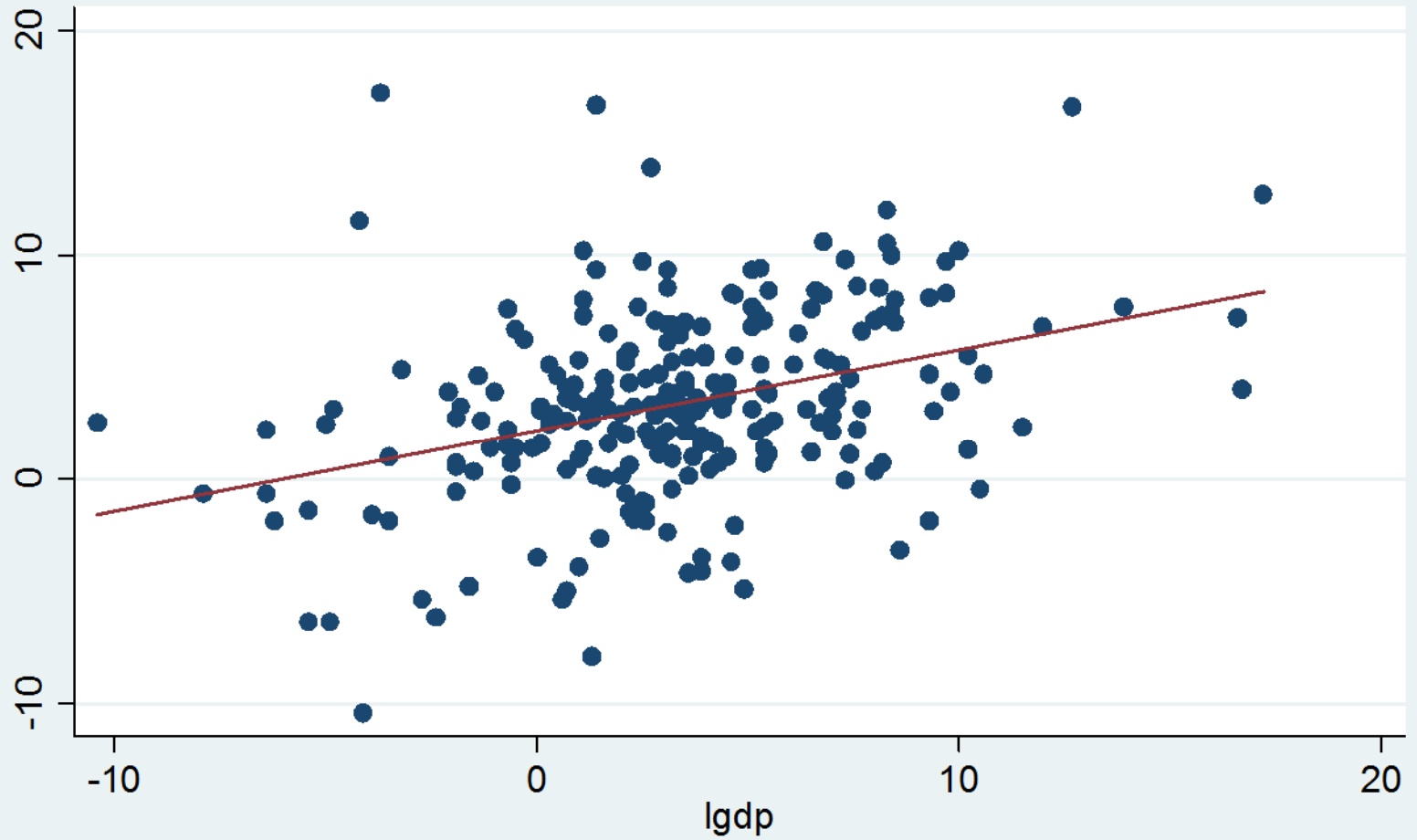
. reg gdp L.gdp

Source	SS	df	MS			
Model	548.684959	1	548.684959	Number of obs = 250		
Residual	3662.76711	248	14.7692222	F( 1, 248) = 37.15		
Total	4211.45207	249	16.9134621	Prob > F = 0.0000		
				R-squared = 0.1303		
				Adj R-squared = 0.1268		
				Root MSE = 3.8431		

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gdp L1.	.3605283	.0591503	6.10	0.000	.2440274	.4770292
_cons	2.146711	.3123873	6.87	0.000	1.531441	2.761982

# GDP Growth Rates



# One-Step-Ahead Forecast

- The optimal forecast for  $T+1$  given  $T$  is

$$\hat{y}_{T+1|T} = \alpha + \beta y_T$$

- The forecast using the estimates is

$$\hat{y}_{T+1|T} = \hat{\alpha} + \hat{\beta} y_T$$

# Example – Unemployment Rate

- The estimates were

<b>ur</b>					
<b>L1.</b>	<b>.9934454</b>	<b>.0051002</b>	<b>194.79</b>	<b>0.000</b>	<b>.9834329</b>
<b>_cons</b>	<b>.045538</b>	<b>.0299153</b>	<b>1.52</b>	<b>0.128</b>	<b>-.0131907</b>

$$y_t = 0.0455 + 0.993y_{t-1} + \hat{e}_t$$

- The value for Jan 2010 is 9.7%, so

$$\hat{y}_{2010:2} = 0.0455 + 0.993 \times 9.7 = 9.68$$

	$y_t$	$y_{t-1}$	fitted value
Jan-09	7.7	7.4	7.40
Feb-09	8.2	7.7	7.70
Mar-09	8.6	8.2	8.19
Apr-09	8.9	8.6	8.59
May-09	9.4	8.9	8.89
Jun-09	9.5	9.4	9.38
Jul-09	9.4	9.5	9.48
Aug-09	9.7	9.4	9.38
Sep-09	9.8	9.7	9.68
Oct-09	10.1	9.8	9.78
Nov-09	10	10.1	10.08
Dec-09	10	10	9.98
Jan-10	9.7	10	9.98
Feb-10	?	9.7	<b>9.68</b>

# Example – GDP Growth

- The estimates were

<b>gdp</b>	<b>Coef.</b>	<b>Std. Err.</b>	<b>t</b>	<b>P&gt; t </b>	<b>[95% Conf. Interval]</b>	
<b>gdp</b> <b>L1.</b>	<b>.3605283</b>	<b>.0591503</b>	<b>6.10</b>	<b>0.000</b>	<b>.2440274</b>	<b>.4770292</b>
<b>_cons</b>	<b>2.146711</b>	<b>.3123873</b>	<b>6.87</b>	<b>0.000</b>	<b>1.531441</b>	<b>2.761982</b>

$$y_t = 2.14 + 0.361y_{t-1} + \hat{e}_t$$

- The value for 4<sup>th</sup> quarter 2009 is 5.7%, so

$$\hat{y}_{2010:1} = 2.14 + 0.361 \times 5.7 = 4.2\%$$

# GDP Growth

	$y_t$	$y_{t-1}$	fitted
2008q4	-5.4	-2.7	1.2
2009q1	-6.4	-5.4	0.2
2009q2	-0.7	-6.4	-0.2
2009q3	2.2	-0.7	1.9
2009q4	5.7	2.2	2.9
2010q1	?	5.7	<b>4.2</b>

# One-Step-Ahead Forecast Error

- The forecast error is

$$\begin{aligned}y_{T+1} - \hat{y}_{T+1|T} &= \alpha + \beta y_T + e_{T+1} - (\alpha + \beta y_T) \\ &= e_{T+1}\end{aligned}$$

- The forecast variance is

$$\text{var}(y_{T+1} - \hat{y}_{T+1|T}) = \text{var}(e_{T+1}) = \sigma^2$$



# Forecast variance estimation

- Average of squared residuals

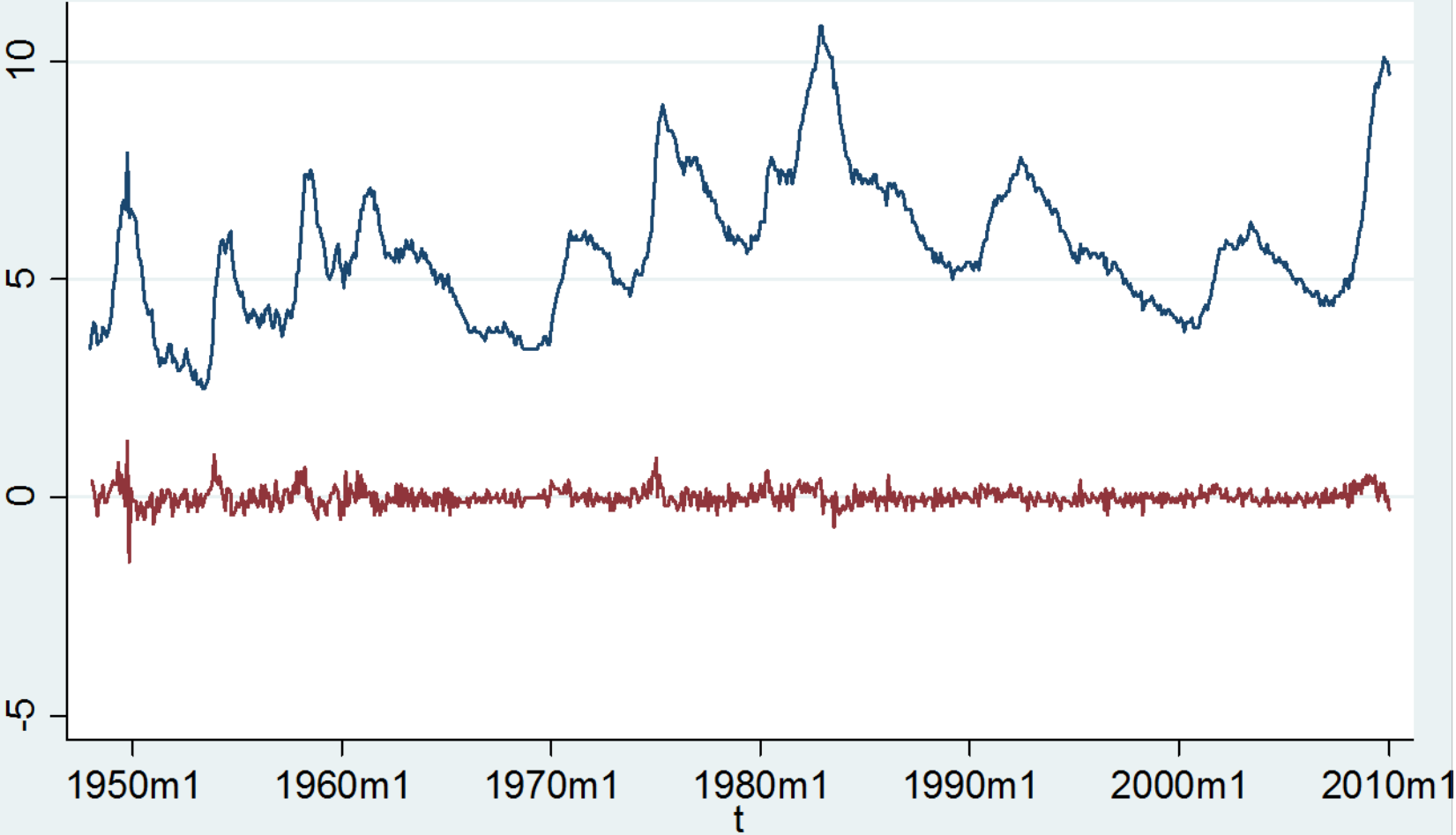
$$\hat{\sigma}^2 = \frac{1}{T} \sum_{t=1}^T \hat{e}_t^2$$

$$\hat{\sigma} = \sqrt{\hat{\sigma}^2}$$

where the least-squares residuals are

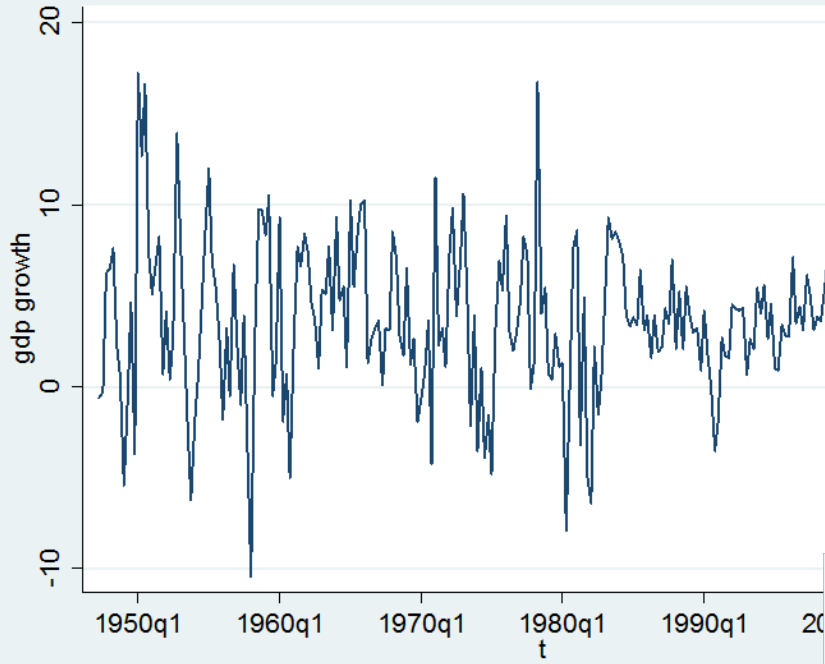
$$\hat{e}_t = y_t - \hat{\alpha} - \hat{\beta}y_{t-1}$$

# Unemployment Rate

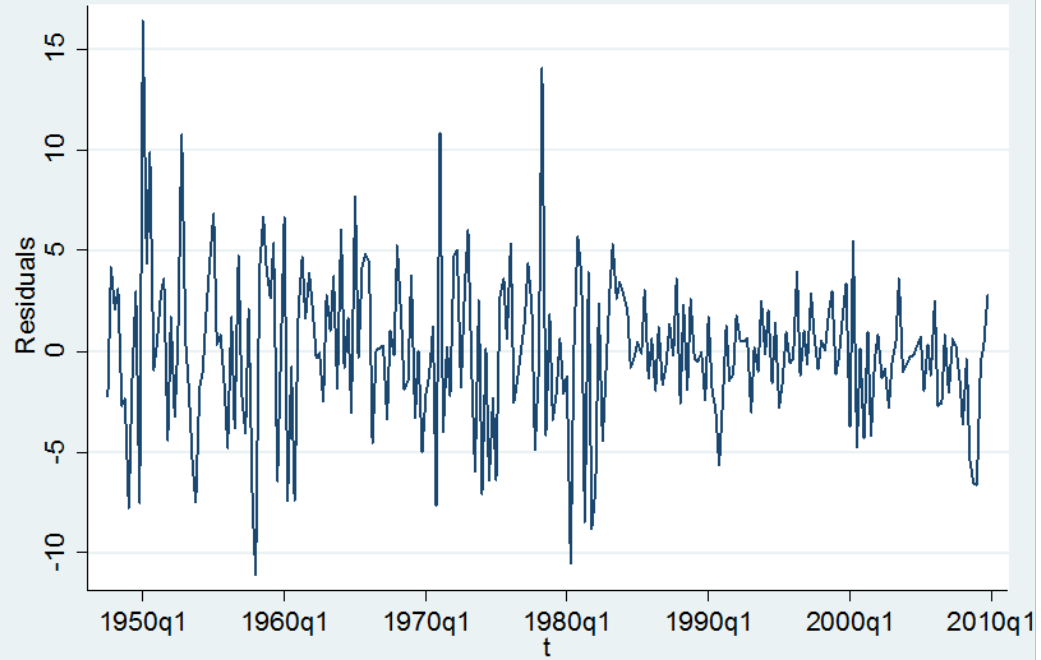


— ur — Residuals

### GDP Growth



### Residuals



# One-Step-Ahead Intervals

- Normal Method

- Assume forecast error is normally distributed
- Forecast interval is point estimate, plus and minus the estimated standard deviation multiplied by a normal quantile

- For a 95% interval:

$$\hat{y}_{T+1|T} \pm \hat{\sigma} \cdot z_{.025} = \hat{y}_{T+1|T} \pm \hat{\sigma} \cdot 1.96$$

- For a 90% interval

$$\hat{y}_{T+1|T} \pm \hat{\sigma} \cdot z_{.05} = \hat{y}_{T+1|T} \pm \hat{\sigma} \cdot 1.645$$

# Estimating Forecast Variance

- The estimated variance is 16.9
- The estimated st. dev. is 3.84

. reg gdp L.gdp

Source	SS	df	MS
Model	548.684959	1	548.684959
Residual	3662.76711	248	14.7692222
Total	4211.45207	249	16.9134621

Number of obs = 250  
 F( 1, 248) = 37.15  
 Prob > F = 0.0000  
 R-squared = 0.1303  
 Adj R-squared = 0.1268  
 Root MSE = 3.8431

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gdp L1.	.3605283	.0591503	6.10	0.000	.2440274	.4770292
_cons	2.146711	.3123873	6.87	0.000	1.531441	2.761982

# Standard Dev Estimation

- In STATA, the estimate of  $\sigma$  or “root mean squared error” is saved after you estimate a regression in “e(rmse)”
- You can access it through the command **e(rmse)**.

```
. dis e(rmse)  
3.8430746
```

# Forecast Interval Construction

```
. tsappend, add(1)

. predict p if t>tq(2009q4)
(option xb assumed; fitted values)
(251 missing values generated)

. gen p1=p-1.645*e(rmse)
(251 missing values generated)

. gen p2=p+1.645*e(rmse)
(251 missing values generated)
```

- Point estimate = 4.2%
- 90% Interval = [-2.1%, 10.5%]

# Two-Step-Ahead Point Forecast

- First Method: Plug-in
- By back-substitution

$$\begin{aligned}y_t &= \alpha + \beta y_{t-1} + e_t \\ &= \alpha + \beta(\alpha + \beta y_{t-2} + e_{t-1}) + e_t \\ &= (1 + \beta)\alpha + \beta^2 y_{t-2} + e_t + \beta e_{t-1}\end{aligned}$$

- Thus

$$\begin{aligned}y_{T+2} &= (1 + \beta)\alpha + \beta^2 y_T + e_{T+2} + \beta e_{T+1} \\ E(y_{T+2} | \Omega_T) &= (1 + \beta)\alpha + \beta^2 y_T\end{aligned}$$



# Plug-in Method

- The optimal forecast is

$$\hat{y}_{T+2|T} = (1 + \beta)\alpha + \beta^2 y_T$$

- Plug-in our estimates to obtain a feasible forecast

$$\hat{y}_{T+2|T} = (1 + \hat{\beta})\hat{\alpha} + \hat{\beta}^2 y_T$$

- This method is feasible but cumbersome for multi-step forecasts and complicated models

# Iterated Method

- Take conditional expectations at time T

$$y_{T+2} = \alpha + \beta y_{T+1} + e_{T+2}$$

$$\begin{aligned} E(y_{T+2} | \Omega_T) &= \alpha + \beta E(y_{T+1} | \Omega_T) + E(e_{T+2} | \Omega_T) \\ &= \alpha + \beta E(y_{T+1} | \Omega_T) \end{aligned}$$

- The left-hand is the 2-step forecast, the right-hand has the 2-step forecast:

$$\hat{y}_{T+2|T} = \alpha + \beta \hat{y}_{T+1|T}$$

# Iterated Method

- We already have the one-step forecast
- The two-step iterates on the one-step

$$\hat{y}_{T+1|T} = \hat{\alpha} + \hat{\beta}y_T$$

$$\hat{y}_{T+2|T} = \hat{\alpha} + \hat{\beta}\hat{y}_{T+1|T}$$

- This method is convenient for linear autoregressions
- It works in **linear** forecasting models (linear autoregressions) but not in nonlinear models
- It is less useful in regression contexts (later sections)

# Direct Method

- We showed that

$$\begin{aligned}y_t &= (1 + \beta)\alpha + \beta^2 y_{t-2} + e_t + \beta e_{t-1} \\ &= \alpha^* + \beta^* y_{t-2} + u_t\end{aligned}$$

where

$$\alpha^* = (1 + \beta)\alpha$$

$$\beta^* = \beta^2$$

$$u_t = e_t + \beta e_{t-1}$$

# Estimation of Direct Method

- This is a regression

$$y_t = \alpha^* + \beta^* y_{t-2} + u_t$$

- The error is the two-step forecast error
- It can be estimated **directly** by least-squares
- This is actually different than the iterated estimator.

# Example – GDP Growth

- $\alpha=2.14$ ,  $\beta=0.361$ ,  $y_T=5.7$ ,  $y_{T+1|T}=4.2$
- Plug-in:

$$\begin{aligned}\hat{y}_{T+2|T} &= (1 + \hat{\beta}) \cdot \hat{\alpha} + \hat{\beta}^2 y_T \\ &= (1 + .36) \cdot 2.14 + .36^2 \cdot 5.7 \\ &= 3.6\%\end{aligned}$$

- Iterated:

$$\begin{aligned}\hat{y}_{T+2|T} &= \hat{\alpha} + \hat{\beta} \hat{y}_{T+1|T} \\ &= 2.14 + .36 \cdot 4.2 \\ &= 3.6\%\end{aligned}$$

# GDP Growth, Direct 2-step

. reg gdp L2.gdp

Source	SS	df	MS			
Model	177.867084	1	177.867084	Number of obs =	249	
Residual	4020.2617	247	16.2763632	F( 1, 247) =	10.93	
Total	4198.12879	248	16.9279387	Prob > F =	0.0011	
				R-squared =	0.0424	
				Adj R-squared =	0.0385	
				Root MSE =	4.0344	

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gdp L2.	.2053005	.0621042	3.31	0.001	.0829792	.3276219
_cons	2.675403	.3285317	8.14	0.000	2.028322	3.322484

- Estimate

$$y_t = 2.67 + 0.21y_{t-2} + \hat{u}_t$$

$$\hat{\sigma} = 4.03$$

- Notice  $.21 > .13 = .36^2$  from iterated

# Direct 2-step-ahead

- 2-step forecast

$$\begin{aligned}\hat{y}_{T+2|T} &= \hat{\alpha}^* + \hat{\beta}^* y_T \\ &= 2.67 + 0.21 \cdot 5.7 \\ &= 3.9\%\end{aligned}$$

- This is slightly larger than the 3.6% forecast from the iterated method



# 2-step forecast variance

- The 2-step equation is

$$y_t = \alpha^* + \beta^* y_{t-2} + u_t$$

$$u_t = e_t + \beta e_{t-1}$$

- The forecast error  $u$  has variance

$$\begin{aligned}\text{var}(u_t) &= \text{var}(e_t + \beta e_{t-1}) \\ &= (1 + \beta^2) \sigma^2\end{aligned}$$

# Plug-in Forecast variance estimation

$$\hat{\sigma}_u^2 = (1 + \hat{\beta}^2) \hat{\sigma}^2$$

$$\hat{\sigma}_u = \sqrt{\hat{\sigma}_u^2}$$

# Iterated Forecast variance estimation

$$\hat{y}_{t|t-1} = \hat{\alpha} + \hat{\beta}y_{t-1}$$

$$\hat{y}_{t|t-2} = \hat{\alpha} + \hat{\beta}\hat{y}_{t|t-1}$$

$$\hat{u}_t = y_t - \hat{y}_{t|t-2}$$

$$\hat{\sigma}_u^2 = \frac{1}{T} \sum_{t=1}^T \hat{u}_t^2$$

$$\hat{\sigma}_u = \sqrt{\hat{\sigma}_u^2}$$

# Direct Forecast variance estimation

$$\hat{u}_t = y_t - \hat{\alpha}^* - \hat{\beta}^* y_{t-2}$$

$$\hat{\sigma}_u^2 = \frac{1}{T} \sum_{t=1}^T \hat{u}_t^2$$

# Example: GDP Growth Plug-in Estimate

- $\beta=.36, \sigma=3.84$

$$\begin{aligned}\hat{\sigma}_u &= \sqrt{(1 + \hat{\beta}^2) \hat{\sigma}^2} \\ &= \sqrt{(1 + .36^2) 3.84^2} \\ &= 4.1\end{aligned}$$

# Iterated Estimate

```
. use gdp
. reg gdp L.gdp
```

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_cons	2.146711	.3123873	6.87	0.000	1.531441	2.761982

```
. predict y1
(option xb assumed; fitted values)
(1 missing value generated)
```

$$\hat{\sigma}_u = 3.95$$

```
. gen y2=_b[_cons]+_b[L.gdp]*y1
(1 missing value generated)
```

```
. gen u=gdp-y2
(1 missing value generated)
```

```
. summarize u
```

variable	obs	Mean	Std. Dev.	Min	Max
u	250	-.0090853	3.951075	-12.78774	14.76027

# Using Estimate Coefficients

- Notice on the previous slide how I used the estimated coefficients in a STATA command
- `gen y2=_b[_cons]+_b[L.gdp]*y1`
- The estimated coefficients are stored in “\_b[]”
- You access them by the name of the variable
  - The constant is always called “\_cons”
    - You can see this in the output
  - The other variable was L.gdp
- The variable “y1” had been created using **predict**

# Direct Estimate

. reg gdp L2.gdp

Source	SS	df	MS			
Model	177.867084	1	177.867084	Number of obs =	249	
Residual	4020.2617	247	16.2763632	F( 1, 247) =	10.93	
Total	4198.12879	248	16.9279387	Prob > F =	0.0011	
				R-squared =	0.0424	
				Adj R-squared =	0.0385	
				Root MSE =	4.0344	

gdp	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
gdp L2.	.2053005	.0621042	3.31	0.001	.0829792	.3276219
_cons	2.675403	.3285317	8.14	0.000	2.028322	3.322484

$$\hat{\sigma} = 4.0$$



# Two-Step-Ahead Intervals

- Normal Method

- Forecast interval is point estimate, plus and minus the estimated standard deviation multiplied by a normal quantile

- For a 95% interval:

$$\hat{y}_{T+2|T} \pm \hat{\sigma}_u \cdot z_{.025} = \hat{y}_{T+2|T} \pm \hat{\sigma}_u \cdot 1.96$$

- For a 90% interval

$$\hat{y}_{T+2|T} \pm \hat{\sigma}_u \cdot z_{.05} = \hat{y}_{T+2|T} \pm \hat{\sigma}_u \cdot 1.645$$

# GDP Growth – Plug-in

- Plug-in Intervals

- $y_{T+2|T} = 3.6\%$ ,  $\sigma_u = 4.1$

- $3.6\% \pm 1.645 * 4.1 = [-3.1\%, 10.3\%]$

- Iterated

- $y_{T+2|T} = 3.6\%$ ,  $\sigma_u = 4.0$

- $3.6\% \pm 1.645 * 4.0 = [-3.0\%, 10.2\%]$

- Direct

- $y_{T+2|T} = 3.9\%$ ,  $\sigma_u = 4.0$

- $3.9\% \pm 1.645 * 4.0 = [-2.7\%, 10.5\%]$

# h-Step-Ahead back substitution

$$\begin{aligned}y_t &= \alpha + \beta y_{t-1} + e_t \\&= \alpha + \beta(\alpha + \beta y_{t-2} + e_{t-1}) + e_t \\&= (1 + \beta)\alpha + \beta^2(\alpha + \beta y_{t-3} + e_{t-2}) + e_t + \beta e_{t-1} \\&= (1 + \beta + \beta^2)\alpha + \beta^3 y_{t-3} + e_t + \beta e_{t-1} + \beta^2 e_{t-2} \\&= (1 + \beta + \beta^2 + \dots + \beta^h)\alpha + \beta^h y_{t-h} + u_t \\u_t &= e_t + \beta e_{t-1} + \beta^2 e_{t-2} + \dots + \beta^{h-1} e_{t-h+1}\end{aligned}$$

# h-Step-Ahead Point Forecast

- Optimal

$$E(y_{T+h} | \Omega_T) = (1 + \beta + \beta^2 + \dots + \beta^h) \alpha + \beta^h y_T$$

- Plug-In

$$\hat{y}_{T+h|T} = (1 + \hat{\beta} + \hat{\beta}^2 + \dots + \hat{\beta}^h) \hat{\alpha} + \hat{\beta}^h y_T$$

- Iterated

$$y_{T+h} = \alpha + \beta y_{T+h-1} + e_{T+h}$$

$$E(y_{T+h} | \Omega_T) = \alpha + \beta E(y_{T+h-1} | \Omega_T)$$

$$\hat{y}_{T+h|T} = \hat{\alpha} + \hat{\beta} \hat{y}_{T+h-1|T}$$

# Direct Method

- Best Linear predictor

$$y_t = \alpha^* + \beta^* y_{t-h} + u_t$$

- Least-Squares estimator

$$y_t = \hat{\alpha}^* + \hat{\beta}^* y_{t-h} + \hat{u}_t$$

- h-step forecast

$$\hat{y}_{T+h|T} = \hat{\alpha}^* + \hat{\beta}^* y_T$$

# Direct Estimates

- Least Squares

$$y_t = 2.14 + 0.36y_{t-1} + \hat{e}_t$$

$$y_t = 2.68 + 0.32y_{t-2} + \hat{u}_t$$

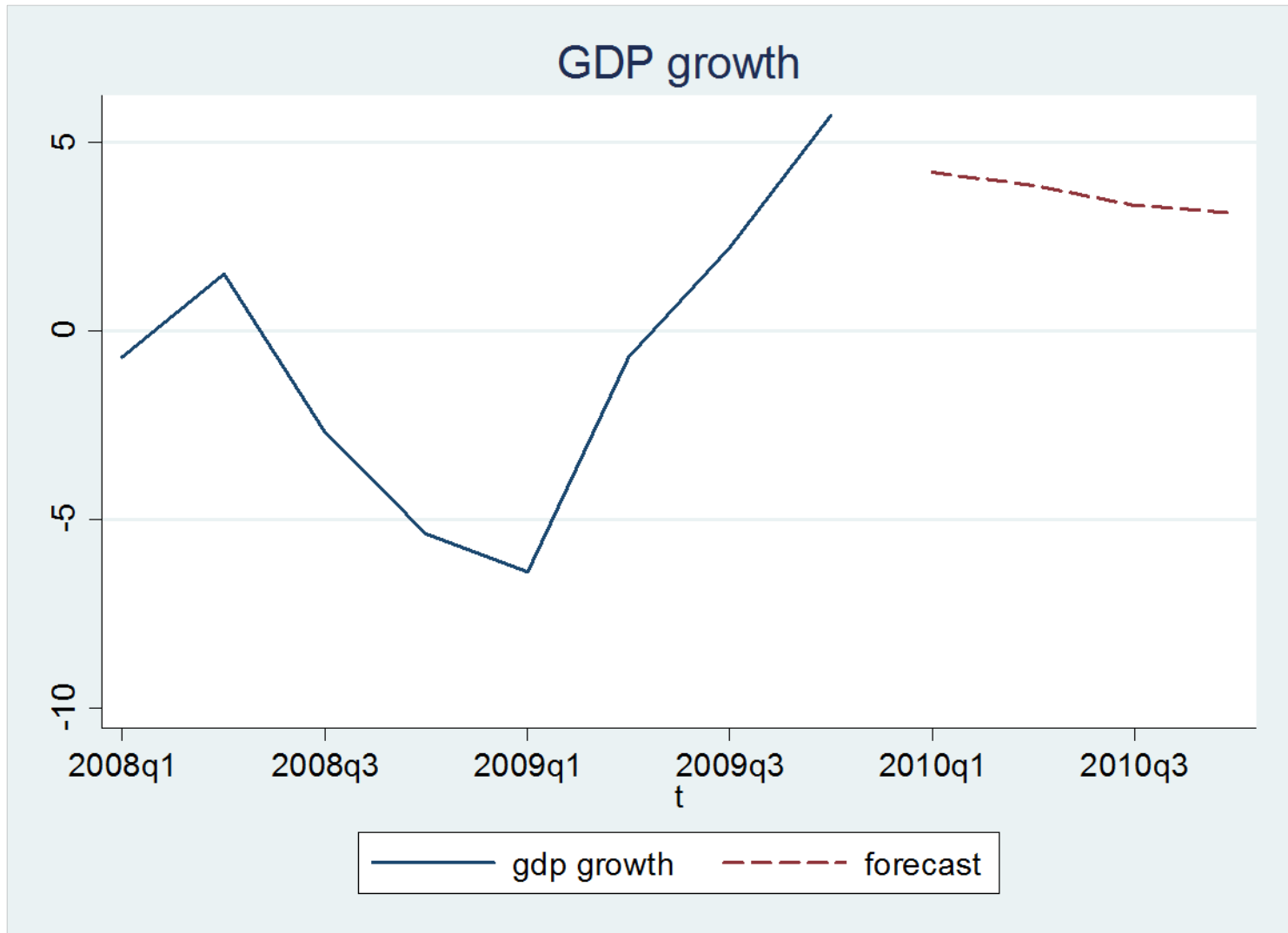
$$y_t = 3.38 - 0.01y_{t-3} + \hat{u}_t$$

$$y_t = 3.65 - 0.09y_{t-4} + \hat{u}_t$$

# Iterated and Direct Point Estimates

	Iterated	Direct
2010Q1	4.2	4.2
2010Q2	3.7	3.8
2010Q3	3.5	3.3
2010Q4	3.4	3.1

# Point Forecast (Direct)





```
use gdp.dta
tsappend, add(4)
reg gdp L.gdp
predict y1
reg gdp L2.gdp
predict y2
reg gdp L3.gdp
predict y3
reg gdp L4.gdp
predict y4
egen p=rowfirst(y1 y2 y3 y4) if t>=tq(2010q1)
label variable p "forecast"
tsline gdp p if t>=tq(2008q1), title(GDP growth) lpattern (solid dash)
```

- There are 4 periods out-of-sample
- The **predict** command computes fitted values for observations which have the needed variables.
- For the regression on the first lag (L.gdp), this works only for the first out-of-sample observation, the remainder are coded as missing.
- For the regression on the second lag (L2.gdp), this works for the first two out-of-sample observations
- etc

# Forecasts

t	y1	y2	y3	y4	p
2009q4	2.94	2.53	3.44	4.16	
2010q1	<b>4.20</b>	3.13	3.39	4.26	<b>4.20</b>
2010q2		<b>3.85</b>	3.36	3.72	<b>3.85</b>
2010q3			<b>3.32</b>	3.44	<b>3.32</b>
2010q4				<b>3.11</b>	<b>3.11</b>

# Multi-Step Forecast Variance

- Can use plug-in, iterated, or direct method
- Easiest method is direct
- Forecast variance is computed from direct regression

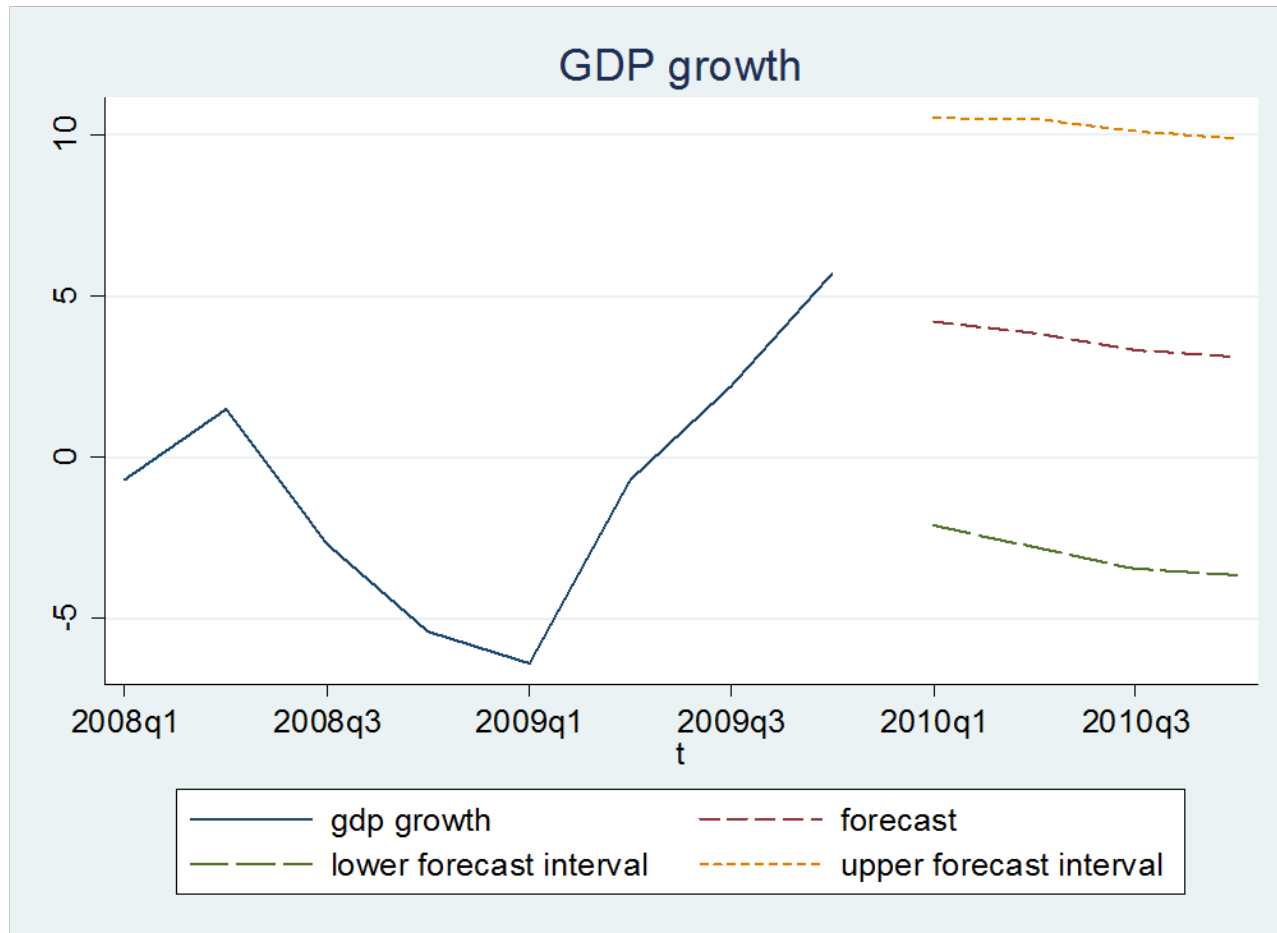
$$y_t = \hat{\alpha}^* + \hat{\beta}^* y_{t-h} + \hat{u}_t$$

$$\hat{\sigma}_u^2 = \frac{1}{T} \sum_{t=1}^T \hat{u}_t^2$$

# Forecast Variance and Intervals

- Regression standard deviation printed in regression output
  - Stored in “e(rmse)”
- Forecast Intervals (90% normal)  
gen y1L=y1-1.645\*e(rmse)  
gen y1U=y1+1.645\*e(rmse)

# GDP Growth Forecast



# .do file

```
tsappend, add(4)
reg gdp L.gdp
predict y1
gen y1L=y1-1.645*e(rmse)
gen y1U=y1+1.645*e(rmse)
reg gdp L2.gdp
predict y2
gen y2L=y2-1.645*e(rmse)
gen y2U=y2+1.645*e(rmse)
reg gdp L3.gdp
predict y3
gen y3L=y3-1.645*e(rmse)
gen y3U=y3+1.645*e(rmse)
reg gdp L4.gdp
predict y4
gen y4L=y4-1.645*e(rmse)
gen y4U=y4+1.645*e(rmse)

egen p=rowfirst(y1 y2 y3 y4) if t>=tq(2010q1)
egen pL=rowfirst(y1L y2L y3L y4L) if t>=tq(2010q1)
egen pU=rowfirst(y1U y2U y3U y4U) if t>=tq(2010q1)
label variable p "forecast"
label variable pL "lower forecast interval"
label variable pU "upper forecast interval"
tsline gdp p pL pU if t>=tq(2008q1), title(GDP growth)
        lpattern (solid dash longdash shortdash)
```